

ON CHARACTERIZATIONS OF THE CONTINUOUS DISTRIBUTIONS BY INDEPENDENCE PROPERTY OF THE QUOTIENT-TYPE UPPER RECORD VALUES

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ABSTRACT. In this paper we obtain characterizations of a family of continuous probability distribution by independence property of upper record values. Also, we introduce some examples of the characterizations of distributions from these general classes of continuous distributions.

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1. Introduction

Record values are important in many real-life situations involving data relating to weather, sports, economics, and life-tests.

Many researchers have studied the characterizations by the independence of the upper record values. For example Ahsanullah [2] studied the characterizations of Pareto distribution by upper record values. Also, Juhás and Skřivánková [5] showed characterization of general classes of distributions with the independent property that the random variables $g(L_n)$ and $g(L_{n+1}) - g(L_n)$. For further various characterizations by the independence of the record values the interested readers are referred to Jin and Lee [4], Lee and Lim [6, 7] among others.

In this paper, we investigate new characterizations of a family of continuous probability distribution by independence property of the quotient-type upper record values.

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2. Main results

In this section, we give our main results. For proving our main theorems we need the following Lemma 2.1. and 2.2..

Lemma 2.1. *Let $\{X_n, n \geq 1\}$ be a sequence of independent and identically distributed random variables with an absolutely continuous cdf $F(x)$ and pdf $f(x)$ on the support (α, β) , where α and β may be finite or infinite. Let $g(x)$ is an increasing and differentiable function with $g(x) \rightarrow 1$ as $x \rightarrow \alpha^+$ and $g(x) \rightarrow \infty$ as $x \rightarrow \beta^-$ for all $x \in (\alpha, \beta)$. Suppose that $\frac{F(g^{-1}(uv))}{F(g^{-1}(u))} = e^{-q(u,v)}$ and $h(u, v) = (q(u, v))^r e^{-q(u,v)} \cdot (\frac{\partial}{\partial v} q(u, v))$, for $r \geq 0$, where $h(u, v) \neq 0$ and $\frac{\partial}{\partial v} q(u, v) \neq 0$ for any positive u and v . If $h(u, v)$ is independent of u , then $q(u, v)$ is a function of v only.*

Proof. Let

$$\begin{aligned} h(u, v) &= (q(u, v))^r e^{-q(u,v)} \left(\frac{\partial}{\partial v} q(u, v) \right) \\ &= \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} (q(u, v))^{k+r} \left(\frac{\partial}{\partial v} q(u, v) \right). \end{aligned} \quad (1)$$

Since $h(u, v)$ is independent of u , we obtain

$$h(u, v) = \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} (q(u, v))^{k+r} \left(\frac{\partial}{\partial v} q(u, v) \right) = l(v). \quad (2)$$

Integrating (2) with respect to v , we get

$$\begin{aligned} \int h(u, v) dv &= \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} \frac{1}{(k+r+1)} (q(u, v))^{k+r+1} \\ &= \int l(v) dv + c = L(v). \end{aligned} \quad (3)$$

Here L is a function of v only and c is independent of v but may depend on u .

Now taking $v \rightarrow 1$, $q(u, v) \rightarrow 0$, we have c independently of u from (3). Differentiating (3) with respect to u , we get

$$\begin{aligned} \frac{\partial}{\partial u} L(v) &= \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} (q(u, v))^{k+r} \left(\frac{\partial}{\partial u} q(u, v) \right) \\ &= l(v) \left(\frac{\partial}{\partial v} q(u, v) \right)^{-1} \left(\frac{\partial}{\partial u} q(u, v) \right) = 0. \end{aligned} \quad (4)$$

Now we know $l(v) \neq 0$ and $\frac{\partial}{\partial v} q(u, v) \neq 0$, so we must have

$$\frac{\partial}{\partial u}q(u, v) = 0. \quad (5)$$

Hence $q(u, v)$ is a function of v only.

This completes the proof. \square

Lemma 2.2. *Let $\{X_n, n \geq 1\}$ be a sequence of independent and identically distributed random variables with an absolutely continuous cdf $F(x)$ and pdf $f(x)$ on the support (α, β) , where α and β may be finite or infinite. Let $g(x)$ is an increasing and differentiable function with $g(x) \rightarrow 0$ as $x \rightarrow \alpha^+$ and $g(x) \rightarrow \infty$ as $x \rightarrow \beta^-$ for all $x \in (\alpha, \beta)$. Suppose that $\frac{R(g^{-1}(uv))}{R(g^{-1}(v))} = q(u, v)$ and $h(u, v) = (1 - q(u, v))^{n-m-1} \times (q(u, v))^{m-1} (\frac{\partial}{\partial u}q(u, v))$ for $1 \leq m < n$, where $R(x) = -\ln \bar{F}(x)$, $h(u, v) \neq 0$ and $\frac{\partial}{\partial u}q(u, v) \neq 0$ for any positive u and v . If $h(u, v)$ is independent of v , then $q(u, v)$ is a function of u only.*

Proof. The proof is similar to the proof of Lemma 2.1.. \square

Theorem 2.3. *Let $\{X_n, n \geq 1\}$ be a sequence of independent and identically distributed random variables each distributed identically with an absolutely continuous cdf $F(x)$ and pdf $f(x)$ on the support (α, β) , where α and β may be finite or infinite. Let $g(x)$ is an increasing and differentiable function with $g(x) \rightarrow 1$ as $x \rightarrow \alpha^+$ and $g(x) \rightarrow \infty$ as $x \rightarrow \beta^-$ for all $x \in (\alpha, \beta)$. Then $F(x) = 1 - (g(x))^{-\lambda}$ for all $g(x) \geq 1, \lambda > 0$, if and only if $\frac{g(X_{U(n)})}{g(X_{U(m)})}$ and $g(X_{U(m)})$ are independent for $1 \leq m < n$.*

Proof. If $F(x) = 1 - (g(x))^{-\lambda}$, then it is easy to see that $\frac{g(X_{U(n)})}{g(X_{U(m)})}$ and $g(X_{U(m)})$ are independent for $1 \leq m < n$.

Let us use the transformation $U = g(X_{U(m)})$ and $V = \frac{g(X_{U(n)})}{g(X_{U(m)})}$. The Jacobian of the transformation is $J = \frac{\partial}{\partial u}(g^{-1}(u)) \times \frac{\partial}{\partial v}(g^{-1}(uv))$. Since $g(x)$ is an increasing and differentiable function, both $\frac{\partial}{\partial u}(g^{-1}(u))$ and $\frac{\partial}{\partial v}(g^{-1}(uv))$ are positive. Thus we can write the joint pdf $f_{U,V}(u, v)$ of U and V as

$$f_{U,V}(u, v) = \frac{1}{\Gamma(m)\Gamma(n-m)} (R(g^{-1}(u)))^{m-1} r(g^{-1}(u)) f(g^{-1}(uv)) \times (R(g^{-1}(uv)) - R(g^{-1}(u)))^{n-m-1} \frac{\partial}{\partial u}(g^{-1}(u)) \frac{\partial}{\partial v}(g^{-1}(uv)) \quad (6)$$

for all $u > 1$ and $v > 1$, where $R(x) = -\ln \bar{F}(x)$ and $r(x) = \frac{d}{dx}R(x)$.

The pdf $f_U(u)$ of U is given by

$$f_U(u) = \frac{R(g^{-1}(u))^{m-1}}{\Gamma(m)} f(g^{-1}(u)) \frac{\partial}{\partial u}(g^{-1}(u)) \quad (7)$$

for all $u > 1$ and $m \geq 1$.

From (6) and (7), we can get the conditional pdf of $f_V(v|u)$ as

$$\begin{aligned} f_V(v|u) &= \frac{(R(g^{-1}(uv)) - R(g^{-1}(u)))^{n-m-1}}{\Gamma(n-m)\bar{F}(g^{-1}(u))} f(g^{-1}(uv)) \frac{\partial}{\partial v}(g^{-1}(uv)) \\ &= \frac{1}{\Gamma(n-m)} \left(-\ln \frac{\bar{F}(g^{-1}(uv))}{\bar{F}(g^{-1}(u))} \right)^{n-m-1} \frac{\bar{F}(g^{-1}(uv))}{\bar{F}(g^{-1}(u))} \left(-\frac{\partial}{\partial v} \left(-\ln \frac{\bar{F}(g^{-1}(uv))}{\bar{F}(g^{-1}(u))} \right) \right), \end{aligned} \quad (8)$$

for all $u > 1$ and $v > 1$.

Since U and V are independent and using Lemma 2.1., $q(u, v) = -\ln \frac{\bar{F}(g^{-1}(uv))}{\bar{F}(g^{-1}(u))}$ is a function of v only. Thus

$$\frac{\bar{F}(g^{-1}(uv))}{\bar{F}(g^{-1}(u))} = L(v) \quad (9)$$

where $L(v)$ is a function of v only. Letting $u \rightarrow 1^+$, so $g^{-1}(u) \rightarrow \alpha$ and $\bar{F}(g^{-1}(u)) \rightarrow 1$, we get $L(v) = \bar{F}(g^{-1}(v))$ from $\bar{F}(\alpha) = 1$. Then, we get

$$\bar{F}(g^{-1}(uv)) = \bar{F}(g^{-1}(u))\bar{F}(g^{-1}(v)) \quad (10)$$

for all $u > 1$ and $v > 1$.

By the theory of functional equation see Aczel(1966), the only continuous solution of (10) with boundary conditions $\bar{F}(\alpha) = 1$ and $\bar{F}(\beta) = 0$ is

$$\bar{F}(x) = (g(x))^{-\lambda} \quad (11)$$

for all $g(x) > 1$ and $\lambda > 0$.

This completes the proof. □

Remark 2.1. If we set $g(x) = x$, $x > 1$, then $F(x) = 1 - x^{-\lambda}$ for all $x > 1$, $\lambda > 0$, if and only if $\frac{X_{U(n)}}{X_{U(m)}}$ and $X_{U(m)}$ are independently distributed in [6].

Remark 2.2. If we set $g(x) = e^x$, $x > 0$, then $F(x) = 1 - e^{-\lambda x}$ for all $x > 0$, $\lambda > 0$, if and only if $X_{U(n)} - X_{U(m)}$ and $X_{U(m)}$ are independently distributed in [3].

Remark 2.3. The case when $m = k$ and $n = k + 1$, which is the special case of Theorem 2.3, have been treated in [4].

Theorem 2.4. Let $\{X_n, n \geq 1\}$ be a sequence of independent and identically distributed random variables each distributed identically with an absolutely continuous cdf $F(x)$ and pdf $f(x)$ on the support (α, β) , where α and β may be finite or infinite. Let $g(x)$ is an increasing and differentiable function with $g(x) \rightarrow 0$ as $x \rightarrow \alpha^+$ and $g(x) \rightarrow \infty$ as $x \rightarrow \beta^-$ for all $x \in (\alpha, \beta)$. Then $F(x) = 1 - e^{-g(x)\lambda}$

for all $g(x) > 0, \lambda > 0$, if and only if $\frac{g(X_{U(m)})}{g(X_{U(n)})}$ and $g(X_{U(n)})$ are independent for $1 \leq m < n$.

Proof. The proof is similar to Theorem 2.3., and is omitted. \square

Remark 2.4. If we set $g(x) = x, x > 0$, then $F(x) = 1 - e^{-x^\lambda}$ for all $x > 0, \lambda > 0$, if and only if $\frac{X_{U(m)}}{X_{U(n)}}$ and $X_{U(n)}$ are independently distributed in [5].

REFERENCES

1. J. Aczel, *Lectures on functional equations and their applications*, Academic Press., NY, 1966.
2. M. Ahsanullah and M. Shakil, *A note on the characterizations of the Pareto distribution by upper record values*, Commun. Korea Math. Soc. **27** (2012), 835-842.
3. M. Ahsanullah and G.G. Hamedani, *Exponential distribution: Theory and Methods*, NOVA Science, New York, 2010.
4. H.W. Jin and M.Y. Lee *On characterization of the continuous distributions by independence property of record values*, J. Appl. Math. and Informatics **35** (2017), 651-657.
5. M. Juhás and V. Skřivánková, *Characterization of general classes of distributions based on independent property of transformed record values*, Applied Mathematics and Computation **226** (2014), 44-50.
6. M.Y. Lee and E.H. Lim, *On characterization of the Pareto distribution by the independent property of record values*, J. Chungcheong Math. Soc. **24** (2011), 85-89.
7. M.Y. Lee and E.H. Lim, *On characterizations of the Weibull distribution by the independent property of record value*, J. Chungcheong Math. Soc. **23** (2010), 245-250.

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