

## A NOTE ON THE CHARACTERIZATIONS OF THE GUMBEL DISTRIBUTION BASED ON LOWER RECORD VALUES

HYUN-WOO JIN\* AND MIN-YOUNG LEE\*\*

ABSTRACT. Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables with cdf  $F(x)$  which is absolutely continuous with pdf  $f(x)$  and  $F(x) < 1$  for all  $x$  in  $(-\infty, \infty)$ .

In this paper, we obtain the characterizations of the Gumbel distribution by lower record values.

### 1. Introduction

Suppose that  $\{X_n, n \geq 1\}$  is a sequence of independent and identically distributed (i.i.d.) random variables with a cumulative distribution function (cdf)  $F(x)$  and a probability density function (pdf)  $f(x)$ . Let  $Y_n = \max(\min)\{X_1, X_2, \dots, X_n\}$  for  $n \geq 1$ . We say  $X_j$  is an upper(lower) record value of  $\{X_n, n \geq 1\}$  if  $Y_j > (<)Y_{j-1}$  for  $j > 1$ . By the definition,  $X_1$  is an upper as well as a lower record value. One can transform the upper records to lower records by replacing the original sequence of  $\{X_j\}$  by  $\{-X_j, j \geq 1\}$  or (if  $P(X_j > 0) = 1$  for all  $j$ ) by  $\{1/X_j, j \geq 1\}$ . The indices at which the upper record values occur are given by record times  $\{U(n), n > 0\}$ , where  $U(n) = \min\{j \mid j > U(n-1), X_j > X_{U(n-1)}\}$ ,  $n > 1$  and  $U(n) = 1$ . We will denote  $L(n)$ ,  $n \geq 1$ , as the indices where lower record values occur.

Shawky and Bakoban (2011) obtained characterization that  $F(x)$  has a Gumbel distribution if and only if  $e^{-L_n} - e^{-L_m}$  and  $e^{L_m}$  are independent for  $1 \leq m < n$ . Also, Nadarajah, Teimouri, and Shih (2014)

---

Received July 20, 2015; Accepted October 17, 2016.

2010 Mathematics Subject Classification: Primary 60E05, 60E10; Secondary 62E50.

Key words and phrases: characterizations, independent, identically distributed, Gumbel distribution, lower record values.

Correspondence should be addressed to Min-Young Lee, [leemy@dankook.ac.kr](mailto:leemy@dankook.ac.kr).

presented characterizations of the Weibull and uniform distributions via the ratio of two record statistics.

In this paper we investigate the characterizations of the Gumbel distributions by lower record values.

## 2. Main results

**THEOREM 2.1.** *Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables with a cdf  $F(x)$  which is absolutely continuous with a pdf  $f(x)$  and  $F(x) < 1$  for all  $x$  in  $(-\infty, \infty)$ . Then  $X_k, k \geq 1$  has a Gumbel distribution if and only if  $X_{L(m)} - X_{L(n)}$  and  $X_{L(n)}$  are independent for  $1 \leq m < n$ .*

*Proof.* If  $X_k \in \text{Gumbel}$ , then it can easily be shown that  $X_{L(n)}$  and  $X_{L(m)} - X_{L(n)}$  are independent for  $1 \leq m < n$ . So we have to prove the converse.

The joint pdf  $f_{m,n}(x, y)$  of  $X_{L(m)}$  and  $X_{L(n)}$  can be written as

$$f_{m,n}(x, y) = \frac{\{H(x)\}^{m-1}}{\Gamma(m)} h(x) \frac{\{H(y) - H(x)\}^{n-m-1}}{\Gamma(n-m)} f(y)$$

where  $H(x) = -\ln F(x)$  and  $h(x) = -\frac{d}{dx} H(x)$ .

Let us use the transformation  $U = X_{L(m)} - X_{L(n)}$  and  $V = X_{L(n)}$ . The Jacobian of the transformation is  $J = 1$ . Thus we can write the joint pdf  $f_{U,V}(u, v)$  of  $U$  and  $V$  as

$$(2.1) \quad \begin{aligned} & f_{U,V}(u, v) \\ &= \frac{\{H(u+v)\}^{m-1}}{\Gamma(m)} h(u+v) \frac{\{H(v) - H(u+v)\}^{n-m-1}}{\Gamma(n-m)} f(v), \end{aligned}$$

for all  $u > 0$  and  $-\infty < v < \infty$ .

Here the marginal pdf  $f_V(v)$  of  $V$  is given by

$$(2.2) \quad f_V(v) = \frac{\{H(v)\}^{n-1}}{\Gamma(n)} f(v), \quad -\infty < v < \infty.$$

From (2.1) and (2.2), we get the conditional pdf of  $f(u | X_{L(n)} = v)$  as

$$\begin{aligned}
 (2.3) \quad & f(u | X_{L(n)} = v) \\
 &= \frac{\Gamma(n)}{\Gamma(m)\Gamma(n-m)} \left(\frac{H(u+v)}{H(v)}\right)^{m-1} \left(1 - \frac{H(u+v)}{H(v)}\right)^{n-m-1} \frac{h(u+v)}{H(v)} \\
 &= \frac{\Gamma(n)}{\Gamma(m)\Gamma(n-m)} \left(\frac{H(u+v)}{H(v)}\right)^{m-1} \left(1 - \frac{H(u+v)}{H(v)}\right)^{n-m-1} \frac{\partial}{\partial u} \left(\frac{H(u+v)}{H(v)}\right)
 \end{aligned}$$

where  $\frac{\partial}{\partial u} \left(\frac{H(u+v)}{H(v)}\right) \neq 0$  for all  $u > 0$  and  $-\infty < v < \infty$ .

Since  $U = X_{L(m)} - X_{L(n)}$  and  $V = X_{L(n)}$  are independent, by using the lemma of Ahasanullah (see p.48 in [2]),  $\frac{\partial}{\partial v} \left(\frac{H(u+v)}{H(v)}\right) = 0$ .

Thus we get

$$(2.4) \quad \frac{H(u+v)}{H(v)} = G(u)$$

where  $G(u)$  is a function of  $u$  only.

By the theory of a functional equation (see [1]), the only continuous solution of (2.4) with the boundary conditions  $H(0) = 1$  and  $H(\infty) = 0$  is

$$H(v) = e^{-v}, \quad -\infty < v < \infty.$$

Thus we have  $F(x) = e^{-e^{-x}}$  for all  $x$  in  $(-\infty, \infty)$ . This completes the proof. □

**THEOREM 2.2.** *Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables with a cdf  $F(x)$  which is absolutely continuous with a pdf  $f(x)$  and  $F(x) < 1$  for all  $x$  in  $(-\infty, \infty)$ . Then  $W = e^{X_{L(m)} - X_{L(n)}}$  has the beta distribution with shape parameter  $n - m$  and  $m$  if and only if  $F(x) = e^{-e^{-x}}$ ,  $-\infty < x < \infty$ .*

*Proof.* The joint pdf  $f_{m,n}(x, y)$  of  $X_{L(m)}$  and  $X_{L(n)}$  can be written as

$$(2.5) \quad f_{m,n}(x, y) = \frac{\{H(x)\}^{m-1}}{\Gamma(m)\Gamma(n-m)} h(x) \{H(y) - H(x)\}^{n-m-1} f(y)$$

for  $1 \leq m < n$ , where  $H(x) = -\ln F(x)$  and  $h(x) = -\frac{d}{dx} H(x)$ .

Consider the functions  $U = X_{L(m)} - X_{L(n)}$  and  $V = X_{L(n)}$ . The Jacobian of the transformation is  $J = 1$ . Thus we can write the joint pdf  $f_{U,V}(u, v)$  of  $U$  and  $V$  as

$$\begin{aligned}
 & f_{U,V}(u, v) \\
 (2.6) \quad &= \frac{\{H(u+v)\}^{m-1}}{\Gamma(m)\Gamma(n-m)} h(u+v)\{H(v) - H(u+v)\}^{n-m-1} f(v)
 \end{aligned}$$

for  $u > 0$  and  $-\infty < v < \infty$ .

If  $F(x) = e^{-e^{-x}}$  for all  $x$  in  $(-\infty, \infty)$ , then we get

$$\begin{aligned}
 & f_{U,V}(u, v) \\
 (2.7) \quad &= \frac{(e^{-(u+v)})^{m-1}}{\Gamma(m)\Gamma(n-m)} e^{-(u+v)} (e^{-v} - e^{-(u+v)})^{n-m-1} e^{-v} e^{-e^{-v}}
 \end{aligned}$$

for  $u > 0$  and  $-\infty < v < \infty$ .

Integrating (2.7) with respect to  $v$ , we have

$$\begin{aligned}
 (2.8) \quad f_U(u) &= \frac{(e^{-u})^m (1 - e^{-u})^{n-m-1}}{\Gamma(m)\Gamma(n-m)} \int_{-\infty}^{\infty} (e^{-v})^n e^{-e^{-v}} dv \\
 &= \frac{\Gamma(n)}{\Gamma(m)\Gamma(n-m)} (e^{-u})^m (1 - e^{-u})^{n-m-1}, \quad u > 0.
 \end{aligned}$$

Now consider the transformation  $W = e^{-U}$ . Then the cdf  $F_W(w)$  of  $W$  is  $F_W(w) = P(e^{-U} \leq w) = P(U \geq -\ln(w)) = 1 - F_U(-\ln(w))$ , and hence the pdf  $f_W(w)$  of  $W$  can be written as

$$(2.9) \quad f_W(w) = \frac{1}{w} f_U(-\ln(w)) = \frac{\Gamma(n)}{\Gamma(m)\Gamma(n-m)} w^{m-1} (1-w)^{n-m-1}$$

for all  $0 < w < 1$ .

Now we will prove the sufficient condition. Let us use the transformation  $U = X_{L(m)} - X_{L(n)}$  and  $V = X_{L(n)}$ . The Jacobian of the transformation is  $J = 1$ . Then we can write the pdf  $f_U(u)$  of  $U$  as

$$\begin{aligned}
 (2.10) \quad f_U(u) &= \int_{-\infty}^{\infty} \frac{\{-\ln F(u+v)\}^{m-1} f(u+v)}{\Gamma(m)\Gamma(n-m)F(u+v)} \\
 &\quad \cdot \{-\ln F(v) + \ln F(u+v)\}^{n-m-1} f(v) dv, \quad u > 0.
 \end{aligned}$$

By setting  $-\ln F(u+v) = t$ , we get

$$(2.11) \quad f_U(u) = \int_0^{\infty} \frac{t^{m-1} \{-\ln F(F^{-1}(e^{-t}) - u) - t\}^{n-m-1}}{\Gamma(m)\Gamma(n-m)} f(F^{-1}(e^{-t}) - u) dt$$

for  $u > 0$ .

Now consider the transformation  $W = e^{-U}$ . Then we get the pdf  $f_W(w)$  of  $W$  as

$$(2.12) \quad \begin{aligned} f_W(w) &= \frac{1}{w} f_U(-\ln(w)) \\ &= \frac{1}{w} \int_0^\infty \frac{t^{m-1} \{-\ln F(F^{-1}(e^{-t}) + \ln(w)) - t\}^{n-m-1}}{\Gamma(m)\Gamma(n-m)} \\ &\quad \cdot f(F^{-1}(e^{-t}) + \ln(w)) dt, \quad 0 < w < 1. \end{aligned}$$

Note that we must have  $F^{-1}(e^{-t}) = -\ln(t)$ , so that the integral in (2.12) must be evaluated to induce a beta pdf with shape parameter  $n - m$  and  $m$ .

Thus we have

$$(2.13) \quad F(x) = e^{-e^{-x}}, \quad -\infty < x < \infty.$$

This completes the proof.  $\square$

## References

- [1] J. Aczel, *Lectures on Functional Equation and Their Applications*, Academic Press, Newyork, 1966.
- [2] M. Ahsanullah, *Record Statistics*, Nova science Publishers, Inc. NJ, USA, 1995.
- [3] S. Nadarajah, M. Teimouri, and S. Shih, *Characterizations of the Weibull and uniform distributions using record values*, Brazilian Journal of Probability and Statistics, **28** (2014), no. 2, 209-222.
- [4] A. I. Shawky & R. A. Bakoban, *Characterization of standard extreme value distributions using records*, J. Chungcheong Math. Soc. **24** (2011), no. 3, 401-407.

\*

Department of Mathematics  
Dankook University  
Cheonan 31116, Republic of Korea  
*E-mail*: hwjin@dankook.ac.kr

\*\*

Department of Mathematics  
Dankook University  
Cheonan 31116, Republic of Korea  
*E-mail*: leemy@dankook.ac.kr