

NEW INFORMATION INEQUALITIES ON ABSOLUTE VALUE OF THE FUNCTIONS AND ITS APPLICATION

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ABSTRACT. Jain and Saraswat (2012) introduced new generalized f - information divergence measure, by which we obtained many well known and new information divergences.

In this work, we introduce new information inequalities in absolute form on this new generalized divergence by considering convex normalized functions. Further, we apply these inequalities for getting new relations among well known divergences, together with numerical verification. Application to the Mutual information is also presented. Asymptotic approximation in terms of Chi- square divergence is done as well.

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1. Introduction

Divergence measures have been demonstrated very useful in a variety of disciplines such as economics and political science [30, 31], biology [23], analysis of contingency tables [10], approximation of probability distributions [5, 18], signal processing [16, 17], pattern recognition [1, 4, 15], color image segmentation [21], 3D image segmentation and word alignment [29], cost- sensitive classification for medical diagnosis [25], magnetic resonance image analysis [32] etc.

Without essential loss of insight, we have restricted ourselves to discrete probability distributions, so let $\Gamma_n = \{P = (p_1, p_2, p_3, \dots, p_n) : p_i > 0, \sum_{i=1}^n p_i = 1\}$, $n \geq 2$ be the set of all complete finite discrete probability distributions. If we take $p_i \geq 0$ for some $i = 1, 2, 3, \dots, n$, then we have to suppose that $0f(0) = 0f\left(\frac{0}{0}\right) = 0$. Some generalized information divergence measures had been introduced, characterized and applied in variety of fields. Such as: Csiszar's f - divergence [6, 7], Bregman's f - divergence [2], Burbea- Rao's f - divergence [3], Renyi's like f -

divergence [24]. Similarly, Jain and Saraswat [14] defined new generalized f -divergence measure, which is given by

$$S_f(P, Q) = \sum_{i=1}^n q_i f\left(\frac{p_i + q_i}{2q_i}\right), \quad (1)$$

where $f : (0, \infty) \rightarrow R$ (set of real no.) is real, continuous, and convex function and $P = (p_1, p_2, \dots, p_n)$, $Q = (q_1, q_2, \dots, q_n) \in \Gamma_n$, where p_i and q_i are probabilities.

Many divergence measures can be obtained from these generalized f -measures by suitably defining the function f . Specially $C_f(P, Q)$ and $S_f(P, Q)$ are widely used due to its compact nature. Some resultant divergences by $S_f(P, Q)$, are as follows.

(a). If we take $f(t) = -\log t$ in (1), we obtain

$$S_f(P, Q) = \sum_{i=1}^n q_i \log\left(\frac{2q_i}{p_i + q_i}\right) = F(Q, P). \quad (2)$$

where $F(Q, P)$ is called adjoint of the Relative JS divergence $F(P, Q)$ [27].

(b). If we take $f(t) = \frac{(t-1)^2}{t}$ in (1), we obtain

$$S_f(P, Q) = \frac{1}{2} \sum_{i=1}^n \frac{(p_i - q_i)^2}{p_i + q_i} = \frac{1}{2} \Delta(P, Q), \quad (3)$$

where $\Delta(P, Q)$ is called the Triangular discrimination [8].

(c). If we take $f(t) = t \log t$ in (1), we obtain

$$S_f(P, Q) = \sum_{i=1}^n \frac{p_i + q_i}{2} \log\left(\frac{p_i + q_i}{2q_i}\right) = G(Q, P), \quad (4)$$

where $G(Q, P)$ is called adjoint of the Relative AG divergence $G(P, Q)$ [28].

(d). If we take $f(t) = (t-1) \log t$ in (1), we obtain

$$S_f(P, Q) = \frac{1}{2} \sum_{i=1}^n (p_i - q_i) \log\left(\frac{p_i + q_i}{2q_i}\right) = \frac{1}{2} J_R(P, Q), \quad (5)$$

where $J_R(P, Q)$ is called the Relative J-divergence [9].

(e). If we take $f(t) = (t-1)^2$ in (1), we obtain

$$S_f(P, Q) = \frac{1}{4} \sum_{i=1}^n \frac{(p_i - q_i)^2}{q_i} = \frac{1}{4} \chi^2(P, Q), \quad (6)$$

where $\chi^2(P, Q)$ is called the Chi-square divergence or Pearson divergence measure [22].

(f). If we take $f(t) = |t-1|$ in (1), we obtain

$$S_f(P, Q) = \frac{1}{2} \sum_{i=1}^n |p_i - q_i| = \frac{1}{2} V(P, Q), \quad (7)$$

where $V(P, Q)$ is called the Variational distance (l_1 distance) [19].
 (g). Particularly, by taking $f(t) = (2t - 1) \log(2t - 1), t \in (\frac{1}{2}, \infty)$ in (1), we obtain

$$S_f(P, Q) = \sum_{i=1}^n p_i \log \frac{p_i}{q_i} = K(P, Q), \tag{8}$$

where $K(P, Q)$ is called the Relative entropy (Kullback- Leibler distance) [20]. Similarly, we can obtain many divergences by using linear convex functions. Since these divergences are not worthful in practice, therefore we can skip them. Now, for a differentiable function $f : (\frac{1}{2}, \infty) \rightarrow R$, consider the associated function $g : (\frac{1}{2}, \infty) \rightarrow R$, which is given by

$$g(t) = (t - 1) f' \left(\frac{t + 1}{2} \right). \tag{9}$$

Put (9) in (1), we get

$$E_{S_f}^*(P, Q) = \sum_{i=1}^n \left(\frac{p_i - q_i}{2} \right) f' \left(\frac{p_i + 3q_i}{4q_i} \right). \tag{10}$$

2. New Information Inequalities

In this section, we introduce new information inequalities on absolute value of the functions for $S_f(P, Q)$. Such inequalities are for instance needed in order to calculate the relative efficiency of two divergences.

Theorem 2.1. *Let f_1 and f_2 be two real, convex and normalized differentiable functions of $(\alpha, \beta) \subset (0, \infty)$, where $0 < \alpha \leq 1 \leq \beta < \infty, \alpha \neq \beta$. If there exists the real constants m, M such that $m < M$, then*

$$m \leq \frac{|f_1(t_1) - f_1(t_2)|}{|f_2(t_1) - f_2(t_2)|} \leq M. \tag{11}$$

By using Cauchy's theorem equation (11) can be written as

$$m \leq \frac{|f_1'(t)|}{|f_2'(t)|} = \left| \frac{f_1'(t)}{f_2'(t)} \right| \leq M, \tag{12}$$

for all $t_1, t_2 \in (\alpha, \beta) \subset (0, \infty)$.

If $P, Q \in \Gamma_n$ is such that $0 < \alpha \leq \frac{p_i + q_i}{2q_i} \leq \beta < \infty \forall i = 1, 2, 3, \dots, n$, then we have the following inequalities

$$mS_{|f_2|}(P, Q) \leq S_{|f_1|}(P, Q) \leq MS_{|f_2|}(P, Q), \tag{13}$$

where $S_f(P, Q)$ is given by (1).

Proof. Firstly, we can see that (12) is obtained from (11) by using Cauchy's theorem.

Now put $t_1 = \frac{p_i + q_i}{2q_i}$ and $t_2 = 1$ in (12), multiply with q_i and then sum over all $i = 1, 2, 3, \dots, n$, we get desire result (13). □

3. Application of New Information Inequalities

In this section, we will obtain bounds of different divergences in terms of the Variational distance by using new inequalities (13). For this, let $f_2 : (0, \infty) \rightarrow R$ be a function defined as

$$f_2(t) = |t - 1|, f_2(1) = 0, f_2'(t) = \begin{cases} -1 & \text{if } 0 < t < 1 \\ 1 & \text{if } 1 < t < \infty \end{cases}, f_2''(t) = 0 \forall t \in (0, \infty)$$

but $f(t)$ is not differentiable at $t = 1$ and

$$|f_2'(t)| = 1. \quad (14)$$

Since $f_2''(t) \geq 0 \forall t > 0 - [1]$ and $f_2(1) = 0$, so $f_2(t)$ is convex and normalized function respectively. Now put $f_2(t)$ in (1), we get

$$S_{|f_2|}(P, Q) = \frac{1}{2} \sum_{i=1}^n |p_i - q_i| = \frac{1}{2} V(P, Q). \quad (15)$$

Proposition 3.1. *Let $V(P, Q)$ and $F(P, Q)$ be defined as in (7) and (2) respectively. For $P, Q \in \Gamma_n$, we have*

$$\frac{1}{2\beta} V(P, Q) \leq |F|(P, Q) \leq \frac{1}{2\alpha} V(P, Q). \quad (16)$$

Proof. Let us consider

$$f_1(t) = -\log t, t \in (0, \infty), f_1(1) = 0, f_1'(t) = -\frac{1}{t} \text{ and } f_1''(t) = \frac{1}{t^2}.$$

Since $f_1''(t) > 0 \forall t > 0$ and $f_1(1) = 0$, so $f_1(t)$ is convex and normalized function respectively.

Now put $f_1(t)$ in (1), we get

$$S_{|f_1|}(P, Q) = \sum_{i=1}^n q_i \left| \log \left(\frac{2q_i}{p_i + q_i} \right) \right| = |F|(Q, P). \quad (17)$$

Now, let $g(t) = \left| \frac{f_1'(t)}{f_2'(t)} \right| = \left| -\frac{1}{t} \right| = \frac{1}{t}$, where $|f_2'(t)| = 1$ and $g'(t) = -\frac{1}{t^2} < 0$.

It is clear that $g(t)$ is always decreasing in $(0, \infty)$, so

$$m = \inf_{t \in (\alpha, \beta)} g(t) = g(\beta) = |f_1'(\beta)| = \frac{1}{\beta}. \quad (18)$$

$$M = \sup_{t \in (\alpha, \beta)} g(t) = g(\alpha) = |f_1'(\alpha)| = \frac{1}{\alpha}. \quad (19)$$

The result (16) is obtained by using (15), (17), (18), and (19) in (13), after interchanging P and Q . \square

Proposition 3.2. Let $V(P, Q)$ and $\Delta(P, Q)$ be defined as in (7) and (3) respectively. For $P, Q \in \Gamma_n$, we have

(a). If $0 < \alpha < 1$, then

$$0 \leq \Delta(P, Q) \leq \frac{1}{2} \left[\frac{\beta^2 - \alpha^2}{\alpha^2 \beta^2} + \left| \frac{\beta^2 + \alpha^2}{\alpha^2 \beta^2} - 2 \right| \right] V(P, Q). \tag{20}$$

(b). If $\alpha = 1$, then

$$0 \leq \Delta(P, Q) \leq \frac{\beta^2 - 1}{\beta^2} V(P, Q). \tag{21}$$

Proof. Let us consider

$$f_1(t) = \frac{(t-1)^2}{t}, t \in (0, \infty), f_1(1) = 0, f_1'(t) = \frac{t^2-1}{t^2} \text{ and } f_1''(t) = \frac{2}{t^3}.$$

Since $f_1''(t) > 0 \forall t > 0$ and $f_1(1) = 0$, so $f_1(t)$ is convex and normalized function respectively.

Now put $f_1(t)$ in (1), we get

$$S_{|f_1|}(P, Q) = \frac{1}{2} \sum_{i=1}^n \frac{(p_i - q_i)^2}{p_i + q_i} = \frac{1}{2} \Delta(P, Q). \tag{22}$$

Now, let

$$g(t) = \left| \frac{f_1'(t)}{f_2'(t)} \right| = \left| \frac{t^2-1}{t^2} \right| = \begin{cases} -\left(\frac{t^2-1}{t^2}\right) & \text{if } 0 < t < 1 \\ \frac{t^2-1}{t^2} & \text{if } 1 \leq t < \infty \end{cases},$$

where $|f_2'(t)| = 1$ and $g'(t) = \begin{cases} -\frac{2}{t^3} < 0 & \text{if } 0 < t < 1 \\ \frac{2}{t^3} > 0 & \text{if } 1 \leq t < \infty \end{cases}.$

It is clear that $g'(t) < 0$ in $(0, 1)$ and > 0 in $(1, \infty)$, i.e., $g(t)$ is decreasing in $(0, 1)$ and increasing in $(1, \infty)$. So $g(t)$ has a minimum value at $t = 1$, therefore

$$m = \inf_{t \in (0, \infty)} g(t) = g(1) = |f_1'(1)| = 0. \tag{23}$$

$$M = \sup_{t \in (\alpha, \beta)} g(t)$$

$$= \begin{cases} \max(|f_1'(\alpha)|, |f_1'(\beta)|) = \frac{|f_1'(\alpha)| + |f_1'(\beta)| + ||f_1'(\alpha)| - |f_1'(\beta)||}{2} & \text{if } 0 < \alpha < 1 \\ |f_1'(\beta)| & \text{if } \alpha = 1 \end{cases},$$

i.e.,

$$M = \begin{cases} \frac{1}{2} \left[\frac{\beta^2 - \alpha^2}{\alpha^2 \beta^2} + \left| \frac{\beta^2 + \alpha^2}{\alpha^2 \beta^2} - 2 \right| \right] & \text{if } 0 < \alpha < 1 \\ \frac{\beta^2 - 1}{\beta^2} & \text{if } \alpha = 1 \end{cases}. \tag{24}$$

The results (20) and (21) are obtained by using (15), (22), (23), and (24) in (13). □

Proposition 3.3. *Let $V(P, Q)$ and $G(P, Q)$ be defined as in (7) and (4) respectively. For $P, Q \in \Gamma_n$, we have*

(a). *If $0 < \alpha \leq \frac{1}{e}$, then*

$$0 \leq |G|(P, Q) \leq \frac{1}{2} \left[\log \sqrt{\frac{\beta}{\alpha}} + \left| \log e\sqrt{\alpha\beta} \right| \right] V(P, Q). \tag{25}$$

(b). *If $\frac{1}{e} < \alpha \leq 1$, then*

$$\frac{\log e\alpha}{2} V(P, Q) \leq |G|(P, Q) \leq \frac{\log e\beta}{2} V(P, Q). \tag{26}$$

Proof. Let us consider

$$f_1(t) = t \log t, t \in (0, \infty), f_1(1) = 0, f_1'(t) = 1 + \log t \text{ and } f_1''(t) = \frac{1}{t}.$$

Since $f_1''(t) > 0 \forall t > 0$ and $f_1(1) = 0$, so $f_1(t)$ is convex and normalized function respectively.

Now put $f_1(t)$ in (1), we get

$$S_{|f_1|}(P, Q) = \sum_{i=1}^n \left(\frac{p_i + q_i}{2} \right) \left| \log \left(\frac{p_i + q_i}{2q_i} \right) \right| = |G|(Q, P). \tag{27}$$

Now, let

$$g(t) = \left| \frac{f_1'(t)}{f_2'(t)} \right| = |1 + \log t| = \begin{cases} -(1 + \log t) & \text{if } 0 < t \leq \frac{1}{e} \\ 1 + \log t & \text{if } \frac{1}{e} < t < \infty \end{cases},$$

where $|f_2'(t)| = 1$ and $g'(t) = \begin{cases} -\frac{1}{t} < 0 & \text{if } 0 < t \leq \frac{1}{e} \\ \frac{1}{t} > 0 & \text{if } \frac{1}{e} < t < \infty \end{cases}.$

It is clear that $g'(t) < 0$ in $(0, \frac{1}{e})$ and > 0 in $(\frac{1}{e}, \infty)$, i.e., $g(t)$ is decreasing in $(0, \frac{1}{e})$ and increasing in $(\frac{1}{e}, \infty)$. So $g(t)$ has a minimum value at $t = \frac{1}{e}$, therefore

$$m = \inf_{t \in (\alpha, \beta)} g(t) = \begin{cases} |f_1'(\frac{1}{e})| = 0 & \text{if } 0 < \alpha \leq \frac{1}{e} \\ |f_1'(\alpha)| = 1 + \log \alpha & \text{if } \frac{1}{e} < \alpha \leq 1 \end{cases}. \tag{28}$$

$$M = \sup_{t \in (\alpha, \beta)} g(t)$$

$$= \begin{cases} \max(|f_1'(\alpha)|, |f_1'(\beta)|) = \left[\log \sqrt{\frac{\beta}{\alpha}} + \left| \log e\sqrt{\alpha\beta} \right| \right] & \text{if } 0 < \alpha \leq \frac{1}{e} \\ |f_1'(\beta)| = 1 + \log \beta & \text{if } \frac{1}{e} < \alpha \leq 1 \end{cases}. \tag{29}$$

The results (25) and (26) are obtained by using (15), (27), (28), and (29) in (13), after interchanging P and Q . □

Proposition 3.4. *Let $V(P, Q)$ and $J_R(P, Q)$ be defined as in (7) and (5) respectively. For $P, Q \in \Gamma_n$, we have*

(a). If $0 < \alpha < 1$, then

$$0 \leq |J_R|(P, Q) \leq \left[\log \sqrt{\frac{\beta}{\alpha}} + \frac{\beta - \alpha}{2\alpha\beta} + \left| \frac{\beta + \alpha}{2\alpha\beta} - \log e\sqrt{\alpha\beta} \right| \right] V(P, Q). \quad (30)$$

(b). If $\alpha = 1$, then

$$0 \leq |J_R|(P, Q) \leq \left(\log e\beta - \frac{1}{\beta} \right) V(P, Q). \quad (31)$$

Proof. Let us consider

$$f_1(t) = (t - 1) \log t, t \in (0, \infty), f_1(1) = 0,$$

$$f_1'(t) = \frac{t - 1}{t} + \log t \text{ and } f_1''(t) = \frac{t + 1}{t^2}.$$

Since $f_1''(t) > 0 \forall t > 0$ and $f_1(1) = 0$, so $f_1(t)$ is convex and normalized function respectively.

Now put $f_1(t)$ in (1), we get

$$S_{|f_1|}(P, Q) = \frac{1}{2} \sum_{i=1}^n \left| (p_i - q_i) \log \left(\frac{p_i + q_i}{2q_i} \right) \right| = \frac{1}{2} |J_R|(P, Q). \quad (32)$$

Now, let

$$g(t) = \left| \frac{f_1'(t)}{f_2'(t)} \right| = \left| \frac{t - 1}{t} + \log t \right| = \begin{cases} -\left(\frac{t-1}{t} + \log t\right) & \text{if } 0 < t < 1 \\ \frac{t-1}{t} + \log t & \text{if } 1 \leq t < \infty \end{cases},$$

where $|f_2'(t)| = 1$ and $g'(t) = \begin{cases} -\left(\frac{t+1}{t^2}\right) < 0 & \text{if } 0 < t < 1 \\ \frac{t+1}{t^2} > 0 & \text{if } 1 \leq t < \infty \end{cases}.$

It is clear that $g'(t) < 0$ in $(0, 1)$ and > 0 in $(1, \infty)$, i.e., $g(t)$ is decreasing in $(0, 1)$ and increasing in $(1, \infty)$. So $g(t)$ has a minimum value at $t = 1$, therefore

$$m = \inf_{t \in (0, \infty)} g(t) = g(1) = |f_1'(1)| = 0. \quad (33)$$

$$M = \sup_{t \in (\alpha, \beta)} g(t)$$

$$= \begin{cases} \max(|f_1'(\alpha)|, |f_1'(\beta)|) = \left[\log \sqrt{\frac{\beta}{\alpha}} + \frac{\beta - \alpha}{2\alpha\beta} + \left| \frac{\beta + \alpha}{2\alpha\beta} - \log e\sqrt{\alpha\beta} \right| \right] & \text{if } 0 < \alpha < 1 \\ \left(\log e\beta - \frac{1}{\beta} \right) & \text{if } \alpha = 1 \end{cases}. \quad (34)$$

The results (30) and (31) are obtained by using (15), (32), (33), and (34) in (13). □

Proposition 3.5. Let $V(P, Q)$ and $\chi^2(P, Q)$ be defined as in (7) and (6) respectively. For $P, Q \in \Gamma_n$, we have

(a). If $0 < \alpha < 1$, then

$$0 \leq \chi^2(P, Q) \leq 2[\beta - \alpha + |2 - (\alpha + \beta)|] V(P, Q). \quad (35)$$

(b). If $\alpha = 1$, then

$$0 \leq \chi^2(P, Q) \leq 4(\beta - 1)V(P, Q). \tag{36}$$

Proof. Let us consider

$$f_1(t) = (t - 1)^2, t \in (0, \infty), f_1(1) = 0, f_1'(t) = 2(t - 1) \text{ and } f_1''(t) = 2.$$

Since $f_1''(t) > 0 \forall t > 0$ and $f_1(1) = 0$, so $f_1(t)$ is convex and normalized function respectively.

Now put $f_1(t)$ in (1), we get

$$S_{|f_1|}(P, Q) = \frac{1}{4} \sum_{i=1}^n \frac{(p_i - q_i)^2}{q_i} = \frac{1}{4} \chi^2(P, Q). \tag{37}$$

Now, let

$$g(t) = \left| \frac{f_1'(t)}{f_2'(t)} \right| = |2(t - 1)| = \begin{cases} -2(t - 1) & \text{if } 0 < t < 1 \\ 2(t - 1) & \text{if } 1 \leq t < \infty \end{cases},$$

where $|f_2'(t)| = 1$ and $g'(t) = \begin{cases} -2 < 0 & \text{if } 0 < t < 1 \\ 2 > 0 & \text{if } 1 \leq t < \infty \end{cases}.$

It is clear that $g'(t) < 0$ in $(0, 1)$ and > 0 in $(1, \infty)$, i.e., $g(t)$ is decreasing in $(0, 1)$ and increasing in $(1, \infty)$. So $g(t)$ has a minimum value at $t = 1$, therefore

$$m = \inf_{t \in (0, \infty)} g(t) = g(1) = |f_1'(1)| = 0. \tag{38}$$

$$\begin{aligned} M &= \sup_{t \in (\alpha, \beta)} g(t) \\ &= \begin{cases} \max(|f_1'(\alpha)|, |f_1'(\beta)|) = [\beta - \alpha + |2 - (\alpha + \beta)|] & \text{if } 0 < \alpha < 1 \\ 2(\beta - 1) & \text{if } \alpha = 1 \end{cases}. \end{aligned} \tag{39}$$

The results (35) and (36) are obtained by using (15), (37), (38), and (39) in (13). \square

4. Numerical Verification

In this section, we give two examples for calculating the divergences $|F|(P, Q), \Delta(P, Q), |J_R|(P, Q)$, and $V(P, Q)$ and then verify the inequalities (16), (20), and (30), numerically.

Example 4.1. Let P be the binomial probability distribution with parameters $(n = 10, p = 0.5)$ and Q its approximated Poisson probability distribution with parameter $(\lambda = np = 5)$ for the random variable X , then we have

x_i	0	1	2	3	4	5	6	7	8	9	10
$p_i \approx$.000976	.00976	.043	.117	.205	.246	.205	.117	.043	.00976	.000976
$q_i \approx$.00673	.033	.084	.140	.175	.175	.146	.104	.065	.036	.018
$\frac{p_i + q_i}{2q_i} \approx$.573	.648	.757	.918	1.086	1.203	1.202	1.063	.831	.636	.527

TABLE 1. Evaluation of probability distributions for $(n = 10, p = 0.5, q = 0.5)$

By using Table 1, we get the followings.

$$\alpha (= .527) \leq \frac{p_i + q_i}{2q_i} \leq \beta (= 1.203). \tag{40}$$

$$|F|(P, Q) = \sum_{i=1}^{11} p_i \left| \log \left(\frac{2p_i}{p_i + q_i} \right) \right| \approx .1495. \tag{41}$$

$$\Delta(P, Q) = \sum_{i=1}^{11} \frac{(p_i - q_i)^2}{p_i + q_i} \approx .0917. \tag{42}$$

$$|J_R|(P, Q) = \sum_{i=1}^{11} \left| (p_i - q_i) \log \left(\frac{p_i + q_i}{2q_i} \right) \right| \approx .0808. \tag{43}$$

$$V(P, Q) = \sum_{i=1}^{11} |p_i - q_i| \approx .3312. \tag{44}$$

Put the approximated numerical values from (40) to (44) in (16), (20), and (30), we obtain the followings respectively.

$$.1376 \leq .1495 (|F|(P, Q)) \leq .3142,$$

$$0 \leq .0917 (\Delta(P, Q)) \leq .8613,$$

and

$$0 \leq .0808 (|J_R|(P, Q)) \leq .51042.$$

Hence verified the inequalities (16), (20), and (30) for $p = 0.5$.

Example 4.2. Let P be the binomial probability distribution with parameters $(n = 10, p = 0.7)$ and Q its approximated Poisson probability distribution with parameter $(\lambda = np = 7)$ for the random variable X , then we have

x_i	0	1	2	3	4	5	6	7	8	9	10
$p_i \approx$.0000059	.000137	.00144	.009	.036	.102	.200	.266	.233	.121	.0282
$q_i \approx$.000911	.00638	.022	.052	.091	.177	.199	.149	.130	.101	.0709
$\frac{p_i + q_i}{2q_i} \approx$.503	.510	.532	.586	.697	.788	1.002	1.392	1.396	1.099	.698

TABLE 2. Evaluation of probability distributions for $(n = 10, p = 0.7, q = 0.3)$

By using Table 2, we get the followings.

$$\alpha (= .503) \leq \frac{p_i + q_i}{2q_i} \leq \beta (= 1.396). \tag{45}$$

$$|F|(P, Q) = \sum_{i=1}^{11} p_i \left| \log \left(\frac{2p_i}{p_i + q_i} \right) \right| \approx .21792. \tag{46}$$

$$\Delta(P, Q) = \sum_{i=1}^{11} \frac{(p_i - q_i)^2}{p_i + q_i} \approx .1812. \tag{47}$$

$$|J_R|(P, Q) = \sum_{i=1}^{11} \left| (p_i - q_i) \log \left(\frac{p_i + q_i}{2q_i} \right) \right| \approx .1686. \quad (48)$$

$$V(P, Q) = \sum_{i=1}^{11} |p_i - q_i| \approx .4844. \quad (49)$$

Put the approximated numerical values from (45) to (49) in (16), (20), and (30), we obtain the followings respectively.

$$.1734 \leq .2179 (|F|(P, Q)) \leq .4815,$$

$$0 \leq .1812 (\Delta(P, Q)) \leq 1.4301,$$

and

$$0 \leq .1686 (|J_R|(P, Q)) \leq .8129.$$

Hence verified the inequalities (16), (20), and (30) for $p = 0.7$.

Remark 4.1. Similarly we can verify the inequalities for different values of p and q and for different discrete probability distributions, like: Geometric, Negative Binomial, Uniform etc.

5. Application to the Mutual Information

Mutual information [26] is a measure of amount of information that one random variable contains about another or amount of information conveyed about one random variable by another.

Let X and Y be two discrete random variables with a joint probability mass function $p(x_i, y_j) = p_{ij}$ with $i = 1, 2, \dots, m; j = 1, 2, \dots, n$ and marginal probability mass functions $p(x_i) = \sum_{j=1}^n p(x_i, y_j), i = 1, 2, \dots, m$ and $p(y_j) = \sum_{i=1}^m p(x_i, y_j), j = 1, 2, \dots, n$, where $x_i \in X, y_j \in Y$, then Mutual information $I(X, Y)$ is defined by

$$I(X, Y) = \sum_{i=1}^m \sum_{j=1}^n p(x_i, y_j) \log \frac{p(x_i, y_j)}{p(x_i)p(y_j)} = \sum_{(x,y) \in (X,Y)} p(x, y) \log \frac{p(x, y)}{p(x)p(y)}.$$

By viewing $K(P, Q)$ (Relative entropy (8)), we can say that the Mutual information is nothing but a Relative entropy between joint distribution $p(x, y)$ and product of marginal distributions $p(x)$ and $p(y)$ after replacing $p(x)$ and $q(x)$ by $p(x, y)$ and $p(x)p(y)$ respectively, in (8). So $I(X, Y)$ can also be written as

$$I(X, Y) = K(p(x, y), p(x)p(y)) = \sum_{(x,y) \in (X,Y)} p(x, y) \log \frac{p(x, y)}{p(x)p(y)}. \quad (50)$$

Similarly, we can define the Mutual information in following manners as well.

In $|F|(P, Q)$ manner:

$$I_{|F|}(X, Y) = \sum_{(x,y) \in (X,Y)} p(x, y) \left| \log \frac{2p(x, y)}{p(x, y) + p(x)p(y)} \right|, \quad (51)$$

In $\chi^2(P, Q)$ manner:

$$I_{\chi^2}(X, Y) = \sum_{(x,y) \in (X,Y)} \frac{[p(x, y) - p(x)p(y)]^2}{p(x)p(y)}, \tag{52}$$

In $J_R(P, Q)$ manner:

$$I_{J_R}(X, Y) = \sum_{(x,y) \in (X,Y)} (p(x, y) - p(x)p(y)) \log \frac{p(x, y) + p(x)p(y)}{2p(x)p(y)}, \tag{53}$$

In $V(P, Q)$ manner:

$$I_V(X, Y) = \sum_{(x,y) \in (X,Y)} |p(x, y) - p(x)p(y)|, \tag{54}$$

and In $|G|(P, Q)$ manner:

$$I_{|G|}(X, Y) = \sum_{(x,y) \in (X,Y)} \left(\frac{p(x, y) + p(x)p(y)}{2} \right) \left| \log \frac{p(x, y) + p(x)p(y)}{2p(x)p(y)} \right|, \tag{55}$$

where $F(P, Q)$, $J_R(P, Q)$, $\chi^2(P, Q)$, $V(P, Q)$, and $G(P, Q)$ are given by (2), (5), (6), (7), and (4) respectively.

So (50) to (55) tell us that how far the joint distribution is from its independency or

$I(X, Y) = 0 = I_{|F|}(X, Y) = I_{\chi^2}(X, Y) = I_{J_R}(X, Y) = I_V(X, Y) = I_{|G|}(X, Y)$ if distributions are independent to each other.

For applications of mutual information, the papers [11] and [12] due to Jain and Chhabra are referred.

Now we introduce the following proposition to obtain results in mutual information sense.

Proposition 5.1. For $\frac{1}{2} < \alpha \leq \frac{p(x,y)+p(x)p(y)}{2p(x)p(y)} \leq \beta < \infty \forall (x, y) \in (X, Y)$, we get the following new information inequalities in Mutual information sense

$$\begin{aligned} |I(X, Y) - I_{J_R}(X, Y)| &\leq \frac{(\beta - \alpha)}{4(2\alpha - 1)(2\beta - 1)} I_{\chi^2}(X, Y) \\ &\leq \frac{(\beta - \alpha)}{2(2\alpha - 1)(2\beta - 1)} (\beta - \alpha + |2 - (\alpha + \beta)|) I_V(X, Y) \\ &\leq \frac{\beta(\beta - \alpha)}{(2\alpha - 1)(2\beta - 1)} (\beta - \alpha + |2 - (\alpha + \beta)|) I_{|F|}(X, Y) \end{aligned} \tag{56}$$

and

$$\begin{aligned} |I(X, Y) - I_{J_R}(X, Y)| &\leq \frac{1}{2} \log \left(\frac{2\beta - 1}{2\alpha - 1} \right) I_V(X, Y) \\ &\leq \frac{1}{1 + \log \alpha} \log \left(\frac{2\beta - 1}{2\alpha - 1} \right) I_{|G|}(X, Y). \end{aligned} \tag{57}$$

Proof. Let us consider

$$f(t) = (2t - 1) \log(2t - 1), t \in \left(\frac{1}{2}, \infty\right), f(1) = 0, f'(t) = 2[1 + \log(2t - 1)]$$

$$\text{and } f''(t) = \frac{4}{2t - 1}. \quad (58)$$

Since $f''(t) > 0 \forall t > \frac{1}{2}$ and $f(1) = 0$, so $f(t)$ is convex and normalized function respectively.

Now put $f(t)$ in (1) and put $f'(t)$ in (10) then replace $p_i, q_i \forall i = 1, 2, \dots, n$ by $p(x, y), p(x)p(y) \forall (x, y) \in (X, Y)$, we get

$$S_f(P, Q) = \sum_{(x, y) \in (X, Y)} p(x, y) \log \frac{p(x, y)}{p(x)p(y)} = I(X, Y) \quad (59)$$

and

$$E_{S_f}^*(P, Q)$$

$$= \sum_{(x, y) \in (X, Y)} (p(x, y) - p(x)p(y)) \log \frac{p(x, y) + p(x)p(y)}{2p(x)p(y)} = I_{J_R}(X, Y) \quad (60)$$

respectively, and

$$A_\alpha^\beta(f') = \int_\alpha^\beta |f''(t)| dt = \int_\alpha^\beta \left| \frac{4}{2t - 1} \right| dt = 2 \log \left(\frac{2\beta - 1}{2\alpha - 1} \right). \quad (61)$$

Now, let $g(t) = f''(t) = \frac{4}{2t - 1}$, where $f''(t)$ is given by (58) and $g'(t) = -\frac{2}{(t - \frac{1}{2})^2} < 0$.

It is clear that $g(t)$ is always decreasing in $(\frac{1}{2}, \infty)$, so

$$m = \inf_{t \in (\alpha, \beta)} g(t) = g(\beta) = \frac{4}{2\beta - 1}. \quad (62)$$

$$M = \sup_{t \in (\alpha, \beta)} g(t) = g(\alpha) = \frac{4}{2\alpha - 1}. \quad (63)$$

The results (56) and (57) are obtained by using (50) to (55), (59) to (63) together with first inequality of (16) and second inequality of (35) in [11] and first inequality of (25) in [12] after replacing p_i, q_i by $p(x, y), p(x)p(y)$ respectively. \square

6. Asymptotic Approximation on $S_f(P, Q)$

In this section, we introduce asymptotic approximation of the generalized f -divergence measure (1) in terms of well known Chi-square divergence (6).

Theorem 6.1. *If $f : (0, \infty) \rightarrow R$ is twice differentiable, convex, and normalized function, i.e., $f''(t) > 0$ and $f(1) = 0$ respectively, then we have*

$$S_f(P, Q) \approx \frac{f''(1)}{8} \chi^2(P, Q). \quad (64)$$

Equivalently

$$\left| \frac{S_f(P, Q)}{\chi^2(P, Q)} - \frac{f''(1)}{8} \right| < \epsilon, \text{ when } |P - Q| < \delta,$$

where $\epsilon, \delta \rightarrow 0$, i.e., ϵ, δ are very small.

Proof. We know by Taylors series expansion of function $f(t)$ at $t = 1$, that

$$f(t) = f(1) + (t - 1)f'(1) + \frac{(t - 1)^2}{2!}f''(1) + (t - 1)^2g(t), \tag{65}$$

where $g(t) = \frac{(t-1)}{3!}f'''(1) + \frac{(t-1)^2}{4!}f''''(1) + \dots$ and we can see that $g(t) \rightarrow 0$ as $t \rightarrow 1$, $f(1) = 0$ because $f(t)$ is normalized, therefore from (65) we get

$$f(t) \approx (t - 1)f'(1) + \frac{(t - 1)^2}{2!}f''(1). \tag{66}$$

Now Put $t = \frac{p_i + q_i}{2q_i}$ in (66), multiply with q_i and then sum over all $i = 1, 2, 3, \dots, n$, we get the desire result (64). \square

7. Conclusion and discussion

In this work, we presented new information inequalities on absolute functions for $S_f(P, Q)$. Further, bounds of various well known divergences have been obtained in terms of the Variational distance in an interval (α, β) , $0 < \alpha \leq 1 \leq \beta < \infty$ with $\alpha \neq \beta$ as an application of new inequalities. These bounds have been verified numerically by taking two discrete distributions: Binomial and Poisson. An approximation on $S_f(P, Q)$ has been done, which relates $S_f(P, Q)$ to $\chi^2(P, Q)$ approximately. Lastly, a very important application to the Mutual information has been discussed, which tells us how far the joint distribution is from its independency.

We found in our previous article [13] that square root of some particular divergences of Csiszar's class is a metric space but not each because of violation of triangle inequality, so we strongly believe that divergence measures can be extended to other significant problems of functional analysis and its applications, such investigations are actually in progress because this is also an area worth being investigated. Such types of divergences are also very useful to find the utility of an event, i.e., an event is how much useful compare to other event.

We hope that this work will motivate the reader to consider the extensions of divergence measures in information theory, other problems of functional analysis and fuzzy mathematics.

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