

A Sufficient Condition on Optimal Berry-Esseen Bounds of Functionals of Gaussian Fields

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Abstract

We find a sufficient condition for optimal Berry-Esseen bounds in the normal approximation of functionals of Gaussian fields studied by Nourdin and Peccati (2009b).

Keywords: Malliavin calculus, Berry-Esseen bound, Stein's method, multiple stochastic integral.

1. Introduction

Let $\{F_n\}$ be a sequence of zero-mean real-valued random variables with the form of a functional of an infinite dimensional Gaussian field. Nourdin and Peccati (2009a), by combining Malliavin calculus with Stein's method, obtain explicit bounds of the type

$$\sup_{z \in R} |P(F_n \leq z) - \Phi(z)| \leq \alpha(n), \quad n \geq 1, \quad (1.1)$$

where $\Phi(z) = \int_{-\infty}^z (1/\sqrt{2\pi})e^{-x^2/2}dx$, and $\alpha(n)$ is some positive sequence converging to zero.

We describe the approach used in Nourdin and Peccati (2009a) as follows. Fix $z \in R$ and consider the Stein equation

$$\mathbf{1}_{(-\infty, z]}(x) - \Phi(x) = f'(x) - xf(x), \quad x \in R. \quad (1.2)$$

Then it is well known that for every $z \in R$, the equation (1.2) has a solution f_z such that $\|f_z\|_\infty \leq \sqrt{2\pi}/4$ and $\|f'_z\|_\infty \leq 1$. Denote by DF_n the Malliavin derivative F_n and by L^{-1} the pseudo-inverse of the Ornstein-Uhlenbeck generator. DF_n is a random element in an appropriate Hilbert space \mathbb{H} . By the integration by parts of a Malliavin calculus,

$$\begin{aligned} P(F_n \leq z) - \Phi(z) &= E \left[f'_z(F_n) - F_n f_z(F_n) \right] \\ &= E \left[f'_z(F_n) \left(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}} \right) \right]. \end{aligned} \quad (1.3)$$

The estimate $\|f'_z\|_\infty \leq 1$ and the Cauchy-Schwartz inequality yield, from (1.3),

$$\sup_{z \in R} |P(F_n \leq z) - \Phi(z)| \leq \sqrt{E \left[(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}})^2 \right]}. \quad (1.4)$$

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Hence the upper bound $\alpha(n)$ appearing in (1.1) is given by

$$\alpha(n) = \sqrt{E[(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}})^2]}.$$

The upper bound $\alpha(n)$ is said to be *optimal* for the sequence $\{F_n\}$ if there exists a constant $c \in (0, 1)$ such that for all sufficiently large n ,

$$c < \frac{\sup_{z \in R} |P(F_n \leq z) - \Phi(z)|}{\alpha(n)} \leq 1. \quad (1.5)$$

Nourdin and Peccati (2009b) show the existence of a constant c appearing in (1.5) by considering the convergence of quantities in (1.3). For this, they assume that the random vectors

$$\left(F_n, \frac{1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}}}{\alpha(n)} \right), \quad n \geq 1, \quad (1.6)$$

converge, in distribution, to a two-dimensional Gaussian vector with non-zero covariance.

Let $\mathbb{D}^{p,q}$ be a class of random variables defined in Section 2. We describe the main result of Nourdin and Peccati (2009b).

Theorem 1. (Nourdin and Peccati) *Let $\{F_n\}$ be a sequence of centered and square integrable functional of some Gaussian process $X = \{X(h) : h \in \mathbb{H}\}$ such that $E(F_n^2) \rightarrow 1$ as $n \rightarrow \infty$. Suppose that the following three conditions hold:*

- (i) *for every n , $F_n \in \mathbb{D}^{1,2}$ and F_n has an absolutely continuous law with respect to Lebesgue measure.*
- (ii) *the quantity $\alpha(n) = \sqrt{E[(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}})^2]}$ is such that (a) $\alpha(n)$ is finite for all n , (b) as $n \rightarrow \infty$, $\alpha(n)$ converges to zero and (c) there exists $m \geq 1$ such that $\alpha(n) > 0$ for $n \geq m$.*
- (iii) *as $n \rightarrow \infty$, the two-dimensional random vectors*

$$\left(F_n, \frac{1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}}}{\alpha(n)} \right)$$

converge, in distribution, to a centered two-dimensional Gaussian vector (N_1, N_2) such that $E(N_1^2) = E(N_2^2) = 1$ and $E(N_1 N_2) = \rho$.

Then, the following upper bound holds:

$$\sup_{z \in R} |P(F_n \leq z) - \Phi(z)| \leq \sqrt{E[(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}})^2]}. \quad (1.7)$$

Moreover, for every $z \in R$,

$$\frac{P(F_n \leq z) - \Phi(z)}{\alpha(n)} \rightarrow \frac{\rho^2}{3}(z^2 - 1) \frac{e^{-\frac{z^2}{2}}}{\sqrt{2\pi}}. \quad (1.8)$$

In this paper, by using Malliavin calculus, we find a sufficient condition on the sequence $\{F_n\}$ to hold the assumption (iii) of Theorem 1.

2. Preliminaries

In this section, we briefly review some basic facts about Malliavin calculus for Gaussian processes. For a more detailed reference, see Nualart (2006). Suppose that \mathbb{H} is a real separable Hilbert space with a scalar product denoted by $\langle \cdot, \cdot \rangle_{\mathbb{H}}$. Let $X = \{X(h), h \in \mathbb{H}\}$ be an isonormal Gaussian process, that is a centered Gaussian family of random variables such that $E(X(h)X(g)) = \langle h, g \rangle_{\mathbb{H}}$. For every $q \geq 1$, let \mathcal{H}_q be the q^{th} Wiener chaos of X , that is the closed linear subspace of $\mathbb{L}^2(\Omega)$ generated by $\{H_q(X(h)) : h \in \mathbb{H}, \|h\|_{\mathbb{H}} = 1\}$, where H_q is the q^{th} Hermite polynomial. We define a linear isometric mapping $I_q : \mathbb{H}^{\otimes q} \rightarrow \mathcal{H}_q$ by $I_q(h^{\otimes q}) = q!H_q(X(h))$, where $\mathbb{H}^{\otimes n}$ is the symmetric tensor product. The following duality formula holds

$$\mathbb{E}[FI_q(h)] = \mathbb{E}[\langle D^q F, h \rangle_{\mathbb{H}^{\otimes q}}], \quad (2.1)$$

for any element $h \in \mathbb{H}^{\otimes q}$ and any random variable $F \in \mathbb{D}^{q,2}$. Here $\mathbb{D}^{q,2}$ is the closure of the set of smooth random variables with respect to the norm

$$\|F\|_{q,2}^2 = \mathbb{E}[F^2] + \sum_{k=1}^q \mathbb{E}[\|D^k F\|_{\mathbb{H}^{\otimes k}}^2],$$

where D^k is the iterative Malliavin derivative.

If $f \in \mathbb{H}^{\otimes p}$, the Malliavin derivative of the multiple stochastic integrals is given by

$$D_z I_q(f_q) = q I_{q-1}(f_q(\cdot, z)), \quad \text{for } z \in [0, 1]^2. \quad (2.2)$$

Let $\{e_l, l \geq 1\}$ be a complete orthonormal system in \mathbb{H} .

The operator L , acting on square integrable random variables, is defined through the projection operators $\{J_q\}_{q \geq 0}$ as $L = \sum_{q=0}^{\infty} -q J_q$, being called the *infinitesimal generator of the Ornstein-Uhlenbeck semigroup*. It has the following properties: F is an element of $\text{Dom}(L) (= \mathbb{D}^{2,2})$ if and only if $F \in \mathbb{D}^{1,2}$ and $DF \in \text{Dom}(\delta)$ and in this case $\delta DF = -LF$. We also define the operator L^{-1} , which is *pseudo-inverse* of L , as follows:

$$L^{-1}F = \sum_{q=1}^{\infty} -\frac{1}{q} J_q(F), \quad \text{for every } F \in \mathbb{L}^2(X).$$

Recall that L^{-1} is an operator with values in $\mathbb{D}^{2,2}$ and that $LL^{-1}F = F - E[F]$ for all $F \in \mathbb{L}^2(X)$.

3. Main Results

We describe our main result in the following Theorem.

Theorem 2. *Let $\{F_n\}$ be a sequence of centered and square integrable functional of some Gaussian process $X = \{X(h) : h \in \mathbb{H}\}$ such that $E(F_n^2) \rightarrow 1$ as $n \rightarrow \infty$ and the condition (i) in Theorem 1 holds. If the following conditions hold,*

(i) *as n tends to infinity,*

$$\|(I - L)^{-\frac{1}{2}}(DF_n)\|_{\mathbb{H}}^2 \rightarrow 1 \quad \text{in } L^2(\Omega).$$

(ii) as n tends to infinity,

$$\frac{1}{\alpha(n)} \left\langle (I - L)^{-1} \left(\langle D^2 F_n, DL^{-1} F_n \rangle_{\mathbb{H}} + \langle DF_n, D^2 L^{-1} F_n \rangle_{\mathbb{H}} \right), DF_n \right\rangle_{\mathbb{H}} \rightarrow \rho \quad \text{in } L^2(\Omega),$$

where

$$\alpha(n) = \sqrt{E[(1 - \langle DF_n, -DL^{-1} F_n \rangle_{\mathbb{H}})^2]}.$$

(iii) as n tends to infinity,

$$\frac{1}{\alpha(n)^2} \left\| (I - L)^{-\frac{1}{2}} \left(\langle D^2 F_n, DL^{-1} F_n \rangle_{\mathbb{H}} + \langle DF_n, D^2 L^{-1} F_n \rangle_{\mathbb{H}} \right) \right\|_{\mathbb{H}}^2 \rightarrow 1 \quad \text{in } L^2(\Omega),$$

then as $n \rightarrow \infty$, the two-dimensional random vectors

$$\left(F_n, \frac{1 - \langle DF_n, -DL^{-1} F_n \rangle_{\mathbb{H}}}{\alpha(n)} \right)$$

converge, in distribution, to a centered two-dimensional Gaussian vector (N_1, N_2) such that $E(N_1^2) = E(N_2^2) = 1$ and $E(N_1 N_2) = \rho$.

Proof: If (N_1, N_2) is a centered two-dimensional Gaussian vector (N_1, N_2) such that $E(N_1^2) = E(N_2^2) = 1$ and $E(N_1 N_2) = \rho$, then the characteristic function of bivariate normal is given by

$$\varphi(s, t) = \exp \left\{ -\frac{1}{2} (s^2 + 2\rho st + t^2) \right\}.$$

We will show that as n tends to infinity,

$$E \left[e^{isF_n + itG_n} \right] \rightarrow \varphi(s, t). \quad (3.1)$$

Let us set

$$G_n = \frac{1 - \langle DF_n, -DL^{-1} F_n \rangle_{\mathbb{H}}}{\alpha(n)}.$$

For every $n \geq 1$, define $\varphi_n(s, t) = E(e^{isF_n + itG_n})$. Then we have

$$\begin{aligned} \frac{\partial}{\partial s} \varphi_n(s, t) &= iE \left[F_n e^{isF_n + itG_n} \right], \\ \frac{\partial}{\partial t} \varphi_n(s, t) &= iE \left[G_n e^{isF_n + itG_n} \right], \\ \frac{\partial^2}{\partial s \partial t} \varphi_n(s, t) &= -E \left[F_n G_n e^{isF_n + itG_n} \right]. \end{aligned}$$

Since $E(F_n^2) \rightarrow 1$ as $n \rightarrow \infty$ and $E(G_n^2) = 1$ for all $n \geq 1$, we have, from Chebyshev's inequality, that for any $\epsilon > 0$ there exists a constant δ_ϵ such that

$$P \left(\sqrt{F_n^2 + G_n^2} \geq \delta_\epsilon \right) \leq \frac{E(F_n^2) + 1}{\delta_\epsilon^2} \leq \frac{C}{\delta_\epsilon^2}.$$

If we take $\delta_\epsilon = \sqrt{\epsilon/C}$, then

$$\sup_n P\left(\sqrt{F_n^2 + G_n^2} \geq \delta_\epsilon\right) \leq \epsilon.$$

Hence the sequence of random vectors (F_n, G_n) is tight. By Prohorov's theorem (Theorem 6.1 in Billingsley (1968)), the sequence (F_n, G_n) is relatively compact, and it suffices to show that the limit of any subsequence converging in distribution is the bivariate normal

$$\mathcal{N}_2\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}\right).$$

We assume that the sequence of random vectors (F_n, G_n) converges, in distribution, to (F, G) , and it suffices to show that

$$(F, G) \sim \mathcal{N}_2\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}\right).$$

Let us set

$$\begin{aligned} \frac{\partial}{\partial s} \varphi(s, t) &= iE\left[F e^{isF+itG}\right], \\ \frac{\partial}{\partial t} \varphi(s, t) &= iE\left[G e^{isF+itG}\right], \\ \frac{\partial^2}{\partial s \partial t} \varphi(s, t) &= -E\left[FG e^{isF+itG}\right]. \end{aligned}$$

It is clear that $G_n e^{isF_n+itG_n}$ converges, in distribution, to $G e^{isF+itG}$. Since $E(G_n^2) = 1$ for all $n \geq 1$, we get, as n goes to infinity,

$$\frac{\partial}{\partial t} \varphi_n(s, t) \rightarrow \frac{\partial}{\partial t} \varphi(s, t).$$

By using the definition of the operator L^{-1} , $L = -\delta D$ and divergence operator δ , we get

$$\begin{aligned} \frac{\partial \varphi_n}{\partial t}(s, t) &= iE\left[L^{-1}L(G_n)e^{isF_n+itG_n}\right] \\ &= -iE\left[\sum_{q \geq 1} \frac{1}{q} J_q(LG_n)e^{isF_n+itG_n}\right] \\ &= -iE\left[\sum_{q \geq 1} \frac{1}{q} L(J_q G_n)e^{isF_n+itG_n}\right] \\ &= i \sum_{q \geq 1} \frac{1}{q} E\left[\delta D(J_q G_n)e^{isF_n+itG_n}\right] \\ &= i \sum_{q \geq 1} \frac{1}{q} E\left[\langle D(J_q G_n), D(e^{isF_n+itG_n}) \rangle_{\mathcal{H}}\right]. \end{aligned}$$

Using the equality $D(J_q F) = J_{q-1}(DF)$ for all $q \geq 1$, we write

$$\begin{aligned} \frac{\partial \varphi_n}{\partial t}(s, t) &= -sE \left[\left\langle \sum_{q=0}^{\infty} \frac{1}{q+1} J_q(DG_n), DF_n \right\rangle_{\mathbb{H}} e^{isF_n+itG_n} \right] - tE \left[\left\langle \sum_{q \geq 1} \frac{1}{q+1} J_q(DG_n), DG_n \right\rangle_{\mathbb{H}} e^{isF_n+itG_n} \right] \\ &= -sE \left[\left\langle (I-L)^{-1}(DG_n), DF_n \right\rangle_{\mathbb{H}} e^{isF_n+itG_n} \right] - tE \left[\left\langle (I-L)^{-1}(DG_n), DG_n \right\rangle_{\mathbb{H}} e^{isF_n+itG_n} \right]. \end{aligned}$$

The above two assumptions (ii) and (iii) give

$$\begin{aligned} \frac{\partial}{\partial t} \varphi_n(s, t) &\rightarrow -s\rho E \left[e^{isF+itG} \right] - tE \left[e^{isF+itG} \right] \\ &= -s\rho \varphi(s, t) - t\varphi(s, t). \end{aligned}$$

Hence we have

$$\frac{\partial \varphi}{\partial t}(s, t) = -(s\rho + t)\varphi(s, t). \quad (3.2)$$

From the differential equation (3.2), we obtain

$$\varphi(s, t) = \exp \left(-\frac{1}{2} (t^2 + 2st\rho) \right) \varphi(s, 0). \quad (3.3)$$

However, similarly as for $(\partial/\partial t)\varphi_n(s, t)$, we get

$$\frac{\partial}{\partial s} \varphi_n(s, 0) = -sE \left[\left\langle (I-L)^{-1}(DF_n), DF_n \right\rangle_{\mathcal{H}} e^{isF_n} \right].$$

The condition (i) yields that as n tends to infinity,

$$\frac{\partial}{\partial s} \varphi_n(s, 0) \rightarrow -s\varphi(s, 0). \quad (3.4)$$

Clearly, $F_n e^{itF_n}$ converges, in distribution, to $F e^{itF}$ and the boundedness of $L^2(\Omega)$ prove that as n tends to infinity,

$$\frac{\partial}{\partial s} \varphi_n(s, 0) \rightarrow \frac{\partial}{\partial s} \varphi(s, 0). \quad (3.5)$$

From (3.4) and (3.5), it follows, with $\varphi(0, 0) = 1$, that

$$\varphi(s, 0) = e^{-\frac{1}{2}s^2}. \quad (3.6)$$

Combining (3.3) with (3.6), we obtain

$$\varphi(s, t) = \exp \left(-\frac{1}{2} (t^2 + 2st\rho + s^2) \right). \quad (3.7)$$

Therefore we complete the proof of (3.1). \square

We give a sufficient condition on $\{F_n\}$ corresponding to (i), (ii) and (iii) in Theorem 2 in the case when $F_n = I_q(f_n)$.

Corollary 1. For $q \geq 2$, consider a sequence $\{F_n = I_q(f_n), n \geq 1\}$ of square integrable random variables belonging to n^{th} Wiener chaos such that

$$E(F_n^2) = q! \|f_n\|_{\mathbb{H}^{\otimes q}}^2 \rightarrow 1 \quad \text{as } n \rightarrow \infty.$$

If the following conditions hold,

(i) as n tends to infinity,

$$q \|I_{q-1}(f_n)\|_{\mathbb{H}}^2 \rightarrow 1 \quad \text{in } L^2(\Omega)$$

(ii) as n tends to infinity,

$$\frac{-2q^2(q-1)}{\alpha(n)} \left\langle (I-L)^{-1} \left(\left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right), I_{q-1}(f_n) \right\rangle_{\mathbb{H}} \rightarrow \rho \quad \text{in } L^2(\Omega),$$

where

$$\alpha(n) = \sqrt{E \left[\left(1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathcal{H}} \right)^2 \right]}.$$

(iii) as n tends to infinity,

$$\frac{4q^2(q-1)^2}{\alpha(n)^2} \left\| (I-L)^{-\frac{1}{2}} \left(\left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right) \right\|_{\mathbb{H}}^2 \rightarrow 1 \quad \text{in } L^2(\Omega),$$

then, as n tends to infinity, the two-dimensional random vectors

$$\left(F_n, \frac{1 - \langle DF_n, -DL^{-1}F_n \rangle_{\mathbb{H}}}{\alpha(n)} \right)$$

converge, in distribution, to a centered two-dimensional Gaussian vector (N_1, N_2) such that $E(N_1^2) = E(N_2^2) = 1$ and $E(N_1 N_2) = \rho$.

Proof: The equations $(I-L)^{-1}(DI_q(f_n)) = I_{q-1}(f_n)$ and $DI_q(f_n) = qI_{q-1}(f_n)$ give

$$\left\langle (I-L)^{-1}(DF_n), DF_n \right\rangle_{\mathbb{H}} = q \|I_{q-1}(f_n)\|_{\mathbb{H}}^2.$$

By using the equations $DL^{-1}(I_q(f_n)) = -I_{q-1}(f_n)$, $D^2 I_q(f_n) = q(q-1)I_{q-2}(f_n)$ and $D^2 L^{-1}(I_q(f_n)) = -(q-1)I_{q-2}(f_n)$, we obtain

$$\left\langle D^2 F_n, DL^{-1}F_n \right\rangle_{\mathbb{H}} = \left\langle DF_n, D^2 L^{-1}F_n \right\rangle_{\mathbb{H}} = -q(q-1) \left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}}. \quad (3.8)$$

From (3.8), it follows that

$$\begin{aligned} & \frac{1}{\alpha(n)} \left\langle (I-L)^{-1} \left(\left\langle D^2 F_n, DL^{-1}F_n \right\rangle_{\mathbb{H}} + \left\langle DF_n, D^2 L^{-1}F_n \right\rangle_{\mathbb{H}} \right), DF_n \right\rangle_{\mathbb{H}} \\ &= \frac{-2q^2(q-1)}{\alpha(n)} \left\langle (I-L)^{-1} \left(\left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right), I_{q-1}(f_n) \right\rangle_{\mathbb{H}}, \end{aligned}$$

which implies the condition (ii). As for the condition (iii), we obtain

$$\begin{aligned}
& \frac{1}{\alpha(n)^2} \left\langle (I - L)^{-1} \left(\left\langle D^2 F_n, DL^{-1} F_n \right\rangle_{\mathbb{H}} + \left\langle DF_n, D^2 L^{-1} F_n \right\rangle_{\mathbb{H}} \right), \left\langle D^2 F_n, DL^{-1} F_n \right\rangle_{\mathbb{H}} + \left\langle DF_n, D^2 L^{-1} F_n \right\rangle_{\mathbb{H}} \right\rangle_{\mathbb{H}} \\
&= \frac{4q^2(q-1)^2}{\alpha(n)^2} \left\langle (I - L)^{-1} \left(\left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right), \left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right\rangle_{\mathbb{H}} \\
&= \frac{4q^2(q-1)^2}{\alpha(n)^2} \left\| (I - L)^{-\frac{1}{2}} \left(\left\langle I_{q-1}(f_n), I_{q-2}(f_n) \right\rangle_{\mathbb{H}} \right) \right\|_{\mathbb{H}}^2.
\end{aligned}$$

Therefore, we complete of the proof of this Corollary. \square

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