

## MARCINKIEWICZ-ZYGMUND LAW OF LARGE NUMBERS FOR BLOCKWISE ADAPTED SEQUENCES

NGUYEN VAN QUANG AND LE VAN THANH

ABSTRACT. In this paper we establish the Marcinkiewicz-Zygmund strong law of large numbers for blockwise adapted sequences. Some related results are considered.

### 1. Introduction and notations

In [5] and [8] it was shown that some properties of independent sequences of random variables can be applied to the sequences consisting of independent blocks. Particularly, it was proved in [8] that if  $(X_i)_{i=1}^\infty$ ,  $EX_i = 0$  is a sequence independent in blocks  $[2^k, 2^{k+1})$ , then it satisfies the Kolmogorov's theorem: the condition  $\sum_{i=1}^\infty (EX_i^2)i^{-2} < \infty$  implies the strong law large numbers (s.l.l.n.), i.e.,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i = 0 \text{ a.s.}$$

Strong law of large numbers for blockwise independent random variables was studied by V. F. Gaposhkin [4].

Marcinkiewicz-Zygmund type strong law of large numbers was studied by many authors. In 1981, N. Etemadi [3] proved that if  $\{X_n, n \geq 1\}$  is a sequence of pairwise i.i.d. random variables with  $E|X_1| < \infty$ , then  $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n (X_i - EX_1) = 0$  a.s.

Later, in 1985, B. D. Choi and S. H. Sung [2] have shown that if  $\{X_n, n \geq 1\}$  are pairwise independent and are dominated in distribution

---

Received January 21, 2005.

2000 Mathematics Subject Classification: 60F05, 60F15.

Key words and phrases: Blockwise independent, blockwise adapted sequence, block martingale difference, Marcinkiewicz-Zygmund law of large numbers.

This work was supported by the National Science Council of Vietnam.

by a random variable  $X$  with  $E|X|^p(\log^+ |X|)^2 < \infty$ ,  $1 < p < 2$ , then

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}}} \sum_{i=1}^n (X_i - EX_i) = 0 \text{ a.s.}$$

Recently, D. H. Hong and S. Y. Hwang [6], D. H. Hong and A. I. Volodin [7] studied Marcinkiewicz-Zygmund strong law of large numbers for double sequence of random variables.

In this paper we establish the Marcinkiewicz-Zygmund strong law of large numbers for blockwise adapted sequences. Some related results are considered.

Let  $\{\omega(n), n \geq 1\}$  be a strictly increasing sequence of positive integers with  $\omega(1) = 1$ . For each  $k \geq 1$ , we set  $\Delta_k = [\omega(k), \omega(k+1))$ . We recall that the sequence  $\{X_i, i \geq 1\}$  of random variables is blockwise independent with respect to blocks  $\Delta_k$ , if for any fixed  $k$ , the sequences  $\{X_i\}_{i \in \Delta_k}$  are independent. Let  $\{\mathcal{F}_i, i \geq 1\}$  be a sequence of  $\sigma$ -fields such that for any fixed  $k$ , the sequences  $\{\mathcal{F}_i, i \in \Delta_k\}$  are increasing. The sequence  $\{X_i, i \geq 1\}$  of random variables is said to be blockwise adapted to  $\{\mathcal{F}_i, i \geq 1\}$ , if each  $X_i$  is measurable with respect to  $\mathcal{F}_i$ . The sequence  $\{X_i, \mathcal{F}_i, i \geq 1\}$  is said to be a block martingale difference with respect to blocks  $\Delta_k$ , if for any fixed  $k$ , the sequences  $\{X_i, \mathcal{F}_i\}_{i \in \Delta_k}$  are martingale differences. Denote

$$\begin{aligned} N_m &= \min\{n | \omega(n) \geq 2^m\}, \\ s_m &= N_{m+1} - N_m + 1, \\ \varphi(i) &= \max_{k \leq m} s_k \text{ if } i \in [2^m, 2^{m+1}), \\ \Delta^{(m)} &= [2^m, 2^{m+1}), m \geq 0, \\ \Delta_k^{(m)} &= \Delta_k \cap \Delta^{(m)}, m \geq 0, k \geq 1, \\ p_m &= \min\{k : \Delta_k^{(m)} \neq \emptyset\}, \\ q_m &= \max\{k : \Delta_k^{(m)} \neq \emptyset\}. \end{aligned}$$

Since  $\omega(N_m - 1) < 2^m$ ,  $\omega(N_m) \geq 2^m$ ,  $\omega(N_{m+1}) \geq 2^{m+1}$  for each  $m \geq 1$ , the number of nonempty blocks  $[\Delta_k^{(m)}]$  is not large than  $s_m = N_{m+1} - N_m + 1$ . Assume  $\Delta_k^{(m)} \neq \emptyset$ , let  $r_k^{(m)} = \min\{r : r \in \Delta_k^{(m)}\}$ .

The sequence  $\{X_n, n \geq 1\}$  is said to be stochastically dominated by a random variable  $X$  if there exists a constant  $C > 0$  such that  $P\{|X_n| > t\} \leq CP\{|X| > t\}$  for all nonnegative real numbers  $t$  and for all  $n > 1$ .

Finally, the symbol  $C$  denotes throughout a generic constant ( $0 < C < \infty$ ) which is not necessarily the same one in each appearance.

**2. Lemmas**

In the sequel we will need the following lemmas.

LEMMA 2.1. (Doob's Inequality) *If  $\{X_i, \mathcal{F}_i\}_{i=1}^N$  is a martingale difference,  $E|X_i|^p < \infty$  ( $1 < p < \infty$ ), then*

$$E \left| \max_{k \leq N} \sum_{i=1}^k X_i \right|^p \leq \left(\frac{p}{p-1}\right)^p E \left| \sum_{i=1}^N X_i \right|^p.$$

The next lemma is due to von Bahr and Esseen [1].

LEMMA 2.2. (von Bahr and Esseen [1]) *Let  $\{X_i\}_{i=1}^N$  be random variables such that  $E\{X_{m+1}|S_m\} = 0$  for  $0 \leq m \leq N-1$ , where  $S_0 = 0$  and  $S_m = \sum_{i=1}^m X_i$  for  $1 \leq m \leq N$ , then*

$$E|S_N|^p \leq C \sum_{i=1}^N E|X_i|^p \text{ for all } 1 \leq p \leq 2,$$

where  $C$  is a constant independent of  $N$ .

By lemmas 2.1 and 2.2, we get the following lemma.

LEMMA 2.3. *If  $\{X_i, \mathcal{F}_i\}_{i=1}^N$  is a martingale difference,  $E|X_i|^p \leq \infty$  ( $1 \leq p \leq 2$ ), then*

$$E \left| \max_{k \leq N} \sum_{i=1}^k X_i \right|^p \leq C \sum_{i=1}^N E|X_i|^p,$$

where  $C$  is a constant independent of  $N$ .

*Proof.* In the case  $p = 1$ , we have

$$E \left| \max_{k \leq N} \sum_{i=1}^k X_i \right| \leq E \left( \sum_{i=1}^N |X_i| \right) = \sum_{i=1}^N E|X_i|.$$

In the case  $1 < p \leq 2$ ,

$$\begin{aligned} E \left| \max_{k \leq N} \sum_{i=1}^k X_i \right|^p &\leq \left(\frac{p}{p-1}\right)^p E \left| \sum_{i=1}^N X_i \right|^p \quad (\text{By Lemma 2.1}) \\ &\leq C \sum_{i=1}^N E|X_i|^p \quad (\text{By Lemma 2.2}). \end{aligned}$$

The proof of the lemma is completed. □

LEMMA 2.4. *If  $q > 1$  and  $\{x_n, n \geq 0\}$  is a sequence of constants such that  $\lim_{n \rightarrow \infty} x_n = 0$ , then*

$$\lim_{n \rightarrow \infty} q^{-n} \sum_{k=0}^n q^{k+1} x_k = 0.$$

*Proof.* Let  $s = q + \sum_{i=0}^{\infty} q^{-i}$ . For any  $\epsilon > 0$ , there exists  $k_0$  such that  $|x_k| < \frac{\epsilon}{2s}$  for all  $k \geq k_0$ . Since  $\lim_{n \rightarrow \infty} q^{-n} = 0$ , so, there exists  $n_0 \geq k_0$  such that  $|q^{-n} \sum_{k=0}^{k_0} q^{k+1} x_k| < \frac{\epsilon}{2}$ . It follows that, for all  $n \geq n_0$ ,

$$\begin{aligned} \left| q^{-n} \sum_{k=0}^n q^{k+1} x_k \right| &\leq \left| q^{-n} \sum_{k=0}^{k_0} q^{k+1} x_k \right| + \left| q^{-n} \sum_{k_0+1}^n q^{k+1} x_k \right| \\ &\leq \frac{\epsilon}{2} + \frac{\epsilon}{2s} (q + 1 + \frac{1}{q} + \dots) \\ &= \frac{\epsilon}{2} + \frac{\epsilon}{2} = \epsilon, \end{aligned}$$

which completes the proof. □

### 3. Main result

With the notations and lemmas accounted for, main results may now be established. Theorem 3.1 establishes the strong law of large numbers for block martingale differences.

THEOREM 3.1. *Let  $\{X_i, \mathcal{F}_i\}_{i=1}^{\infty}$  be a block martingale difference with respect to blocks  $\Delta_k$ , ( $1 \leq p \leq 2$ ). If*

$$\sum_{i=1}^{\infty} \frac{E|X_i|^2}{i^{\frac{2}{p}}} < \infty,$$

then

$$\frac{\sum_{i=1}^n X_i}{n^{\frac{1}{p}} \varphi^{\frac{1}{2}}(n)} \rightarrow 0 \quad \text{a.s. as } n \rightarrow \infty.$$

*Proof.* Let

$$\gamma_k^{(m)} = \max_{n \in \Delta_k^{(m)}} \left| \sum_{i=r_k^{(m)}}^n X_i \right|, \quad m \geq 0, k \geq 1;$$

$$\gamma_m = 2^{-\frac{m-1}{p}} \varphi^{-\frac{1}{2}}(2^m) \sum_{p_m \leq k \leq q_m} \gamma_k^{(m)}, \quad m \geq 0.$$

Using Lemma 2.3 for martingale differences  $\{X_i, \mathcal{F}_i, i \in \Delta_k^{(m)}\}$ , we have

$$E|\gamma_k^{(m)}|^2 \leq C \sum_{i \in \Delta_k^{(m)}} EX_i^2, \text{ for all } m \geq 0, k \geq 1.$$

It implies

$$\begin{aligned} E|\gamma_m|^2 &\leq 2^{-\frac{2m-2}{p}} \varphi^{-1}(2^m) s_m \sum_{k=p_m}^{q_m} E|\gamma_k^{(m)}|^2 \\ &\leq 2^{-\frac{2m-2}{p}} \sum_{k=p_m}^{q_m} E|\gamma_k^{(m)}|^2 \\ &\leq C 2^{-\frac{2m-2}{p}} \sum_{i=2^m}^{2^{m+1}-1} EX_i^2 \\ &\leq C \sum_{i=2^m}^{2^{m+1}-1} \frac{X_i^2}{i^{\frac{2}{p}}}. \end{aligned}$$

Thus

$$\sum_{m=0}^{\infty} E|\gamma_m|^2 \leq C \sum_{i=1}^{\infty} \frac{X_i^2}{i^{\frac{2}{p}}} < \infty.$$

By the Markov inequality and the Borel-Cantelli lemma, we get

$$(3.1) \quad \lim_{m \rightarrow \infty} \gamma_m = 0 \quad a.s.$$

On the other hand

$$(3.2) \quad 0 \leq 2^{-\frac{m}{p}} \varphi^{-\frac{1}{2}}(2^m) \sum_{k=0}^m \sum_{i=p_k}^{q_k} \gamma_i^{(k)} \leq 2^{-\frac{m}{p}} \sum_{k=0}^m 2^{\frac{k+1}{p}} \gamma_k.$$

By (3.1), (3.2) and Lemma 2.4, we get

$$(3.3) \quad \lim_{m \rightarrow \infty} 2^{-\frac{m}{p}} \varphi^{-\frac{1}{2}}(2^m) \sum_{k=0}^m \sum_{i=p_k}^{q_k} \gamma_i^{(k)} = 0 \quad a.s.$$

Assume  $n \in \Delta_k^{(m)}$ , we have

$$\begin{aligned}
 0 &\leq \left| n^{-\frac{1}{p}} \varphi^{-\frac{1}{2}}(n) \sum_{i=1}^n X_i \right| \\
 (3.4) \quad &\leq 2^{-\frac{m}{p}} \varphi^{-\frac{1}{2}}(2^m) \sum_{k=0}^m \sum_{i=p_k}^{q_k} \gamma_i^{(k)}.
 \end{aligned}$$

By (3.3) and (3.4), we get

$$n^{-\frac{1}{p}} \varphi^{-\frac{1}{2}}(n) \sum_{i=1}^n X_i \rightarrow 0 \text{ a.s. (as } n \rightarrow \infty).$$

The proof is completed. □

In the next theorem, we set up the Marcinkiewicz-Zygmund law of large numbers for blockwise adapted sequences which are stochastically dominated by a random variable  $X$ .

**THEOREM 3.2.** *Let  $\{\mathcal{F}_i, i \geq 1\}$  be a sequence of  $\sigma$ -fields such that for any fixed  $k$ , the sequences  $\{\mathcal{F}_i, i \in \Delta_k\}$  are increasing and  $\{X_i, i \geq 1\}$  is blockwise adapted to  $\{\mathcal{F}_i, i \geq 1\}$ . If  $\{X_i, i \geq 1\}$  is stochastically dominated by a random variable  $X$  such that either*

$$E|X| \log^+ |X| < \infty \text{ if } p = 1,$$

or

$$E|X|^p < \infty \text{ if } 1 < p < 2,$$

then

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}} \varphi^{\frac{1}{2}}(n)} \sum_{i=1}^n (X_i - a_i) = 0 \text{ a.s.},$$

where  $a_i = EX_i$  if  $i = r_k^{(m)}$  and  $a_i = E(X_i | \mathcal{F}_{i-1})$  if  $i \neq r_k^{(m)}$  for  $k \geq 1$  and  $m \geq 0$ .

*Proof.* Let  $X'_i = X_i I\{|X_i| \leq i^{\frac{1}{p}}\}$ ,  $b_i = EX'_i$  if  $i = r_k^{(m)}$  and  $b_i = E(X'_i | \mathcal{F}_{i-1})$  if  $i \neq r_k^{(m)}$  for  $k \geq 1$  and  $m \geq 0$ . We have

$$\begin{aligned}
 E(X'_i - b_i)^2 &\leq E|X'_i|^2 \\
 &= \int_0^{i^{\frac{2}{p}}} P(|X_i|^2 > t) dt
 \end{aligned}$$

$$\begin{aligned} &\leq C \int_0^{i^{\frac{2}{p}}} P(|X|^2 > t) dt \\ &= C \int_0^{i^{\frac{2}{p}}} (P(t < |X|^2 < i^{\frac{2}{p}}) + P(i^{\frac{2}{p}} \leq |X|^2)) dt \\ &= C \left( \int_0^{i^{\frac{1}{p}}} x^2 dF(x) + i^{\frac{2}{p}} P(i^{\frac{2}{p}} \leq |X|^2) \right), \end{aligned}$$

where  $F(x)$  is the distribution function of  $X$ .

$$\begin{aligned} \sum_{i=1}^{\infty} \frac{1}{i^{\frac{2}{p}}} \int_0^{i^{\frac{1}{p}}} x^2 dF(x) &\leq C \sum_{i=1}^{\infty} \frac{1}{i^{\frac{2}{p}}} \sum_{k=1}^i \int_{(k-1)^{\frac{1}{p}}}^{k^{\frac{1}{p}}} x^2 dF(x) \\ &\leq C \sum_{k=1}^{\infty} \sum_{i=k}^{\infty} \frac{1}{i^{\frac{2}{p}}} \int_{(k-1)^{\frac{1}{p}}}^{k^{\frac{1}{p}}} x^2 dF(x) \\ &\leq C \sum_{k=1}^{\infty} k^{\frac{p-2}{p}} \int_{(k-1)^{\frac{1}{p}}}^{k^{\frac{1}{p}}} x^2 dF(x) \\ &\leq C \sum_{k=1}^{\infty} \int_{(k-1)^{\frac{1}{p}}}^{k^{\frac{1}{p}}} x^p dF(x) \\ &\leq CE|X|^p < \infty, \end{aligned}$$

and

$$\sum_{i=1}^{\infty} P(i^{\frac{2}{p}} \leq |X|^2) = \sum_{i=1}^{\infty} P(i \leq |X|^p) \leq CE|X|^p < \infty.$$

Hence

$$\sum_{i=1}^{\infty} \frac{E(X'_i - b_i)^2}{i^{\frac{2}{p}}} < \infty.$$

For each  $k > 1$  and  $m \geq 0$ , sequence  $\{X'_i - b_i, \mathcal{F}_i, i \in \Delta_k^{(m)}\}$  is a martingale difference. By using the proof of Theorem 3.1, we get

$$(3.5) \quad n^{-\frac{1}{p}} \varphi^{-\frac{1}{2}}(n) \sum_{i=1}^n (X'_i - b_i) \rightarrow 0 \text{ a.s. (as } n \rightarrow \infty).$$

Next,

$$\begin{aligned}
 \sum_{i=1}^{\infty} P(X_i \neq X'_i) &= \sum_{i=1}^{\infty} P(|X_i| > i^{\frac{1}{p}}) \\
 &\leq C \sum_{i=1}^{\infty} P(|X| > i^{\frac{1}{p}}) \\
 (3.6) \quad &\leq C \sum_{i=1}^{\infty} P(|X|^p > i) \leq CE|X|^p < \infty.
 \end{aligned}$$

Finally, we prove that

$$(3.7) \quad \lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}}} \sum_{i=1}^n (a_i - b_i) = 0 \text{ a.s.}$$

In the case  $p=1$ ,

$$\begin{aligned}
 \sum_{n=1}^{\infty} n^{-1} E[|X_n| I(|X_n| > n)] &= \sum_{n=1}^{\infty} n^{-1} \int_n^{\infty} P(|X_n| > x) dx \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} \int_n^{\infty} P(|X| > x) dx \\
 &= C \sum_{n=1}^{\infty} n^{-1} \sum_{i=n}^{\infty} \int_{i < x \leq i+1} P(|X| > x) dx \\
 &\leq C \sum_{i=1}^{\infty} P(|X| > i) \sum_{n=1}^i n^{-1} \\
 &\leq C \sum_{i=1}^{\infty} (1 + \log i) P(|X| > i) < \infty.
 \end{aligned}$$

This implies that  $\sum_{n=1}^{\infty} n^{-1} (a_n - b_n) < \infty$  a.s. By using Kronecker's lemma, we get (3.7).

In the case  $1 < p < 2$ , since

$$\sum_{n=1}^{\infty} n^{-\frac{1}{p}} E[|X_n| I(|X_n| > n^{\frac{1}{p}})]$$



$$\begin{aligned}
 &\leq C \sum_{n=1}^{\infty} n^{-\frac{1}{p}} \int_{\frac{1}{n^{\frac{1}{p}}}}^{\infty} x dF(x) \\
 &= C \sum_{n=1}^{\infty} n^{-\frac{1}{p}} \sum_{i=n}^{\infty} \int_{\frac{1}{i^{\frac{1}{p}}}}^{(i+1)^{\frac{1}{p}}} x dF(x) \\
 &\leq C \sum_{i=1}^{\infty} \sum_{n=1}^i n^{-\frac{1}{p}} \int_{\frac{1}{i^{\frac{1}{p}}}}^{(i+1)^{\frac{1}{p}}} x dF(x) \\
 &\leq C \sum_{i=1}^{\infty} i^{\frac{p-1}{p}} \int_{\frac{1}{i^{\frac{1}{p}}}}^{(i+1)^{\frac{1}{p}}} x dF(x) \\
 &\leq C \sum_{i=1}^{\infty} \int_{\frac{1}{i^{\frac{1}{p}}}}^{(i+1)^{\frac{1}{p}}} x^p dF(x) < \infty.
 \end{aligned}$$

It follows that

$$\sum_{n=1}^{\infty} n^{-\frac{1}{p}} (a_n - b_n) < \infty \text{ a.s.}$$

By Kronecker’s lemma, we get (3.7).

Combining (3.5), (3.6) and (3.7) we obtain

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}} \varphi^{\frac{1}{2}}(n)} \sum_{i=1}^n (X_i - a_i) = 0 \text{ a.s.}$$

This completes the proof of theorem. □

The following corollaries extend the classical Marcinkiewicz-Zygmund strong law of large numbers.

**COROLLARY 3.3.** *Let  $\{X_i, i \geq 1\}$  be a sequence of blockwise independent random variables with respect to blocks  $\Delta_k$ . If  $\{X_i, i \geq 1\}$  is stochastically dominated by a random variable  $X$ ,  $E|X|^p < \infty$  ( $1 \leq p < 2$ ), then*

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}} \varphi^{\frac{1}{2}}(n)} \sum_{i=1}^n (X_i - EX_i) = 0 \text{ a.s.}$$

*Proof.* Let  $\mathcal{F}_i = \sigma(X_{r_k^{(m)}}, \dots, X_i)$  (the  $\sigma$ -field generated by  $X_{r_k^{(m)}}, \dots, X_i$ ) if  $i \in \Delta_k^{(m)}$ . Then  $\{X_i, i \geq 1\}$  is blockwise adapted to  $\{\mathcal{F}_i, i \geq 1\}$ . From the independence of sequence  $\{X_i, i \in \Delta_k^{(m)}\}$  we get for all  $k$  and  $m$

$$E(X_i | \mathcal{F}_{i-1}) = EX_i \text{ if } i \neq r_k^{(m)}.$$

By the proof of Theorem 3.2, we only need prove for the case  $p = 1$ .

In the case  $p = 1$ , also using the proof of Theorem 3.2, we get

$$(3.8) \quad n^{-1}\varphi^{-\frac{1}{2}}(n) \sum_{i=1}^n (X_i - EX'_i) \rightarrow 0 \text{ (as } n \rightarrow \infty),$$

where  $X'_i = X_i I(|X_i| \leq i)$ . On the other hand

$$\begin{aligned} E[|X_i|I(|X_i| > i)] &= \int_i^\infty P(|X_i| > x)dx \\ &\leq C \int_i^\infty P(|X| > x)dx \rightarrow 0 \text{ as } i \rightarrow \infty. \end{aligned}$$

Thus

$$(3.9) \quad |n^{-1} \sum_{i=1}^n (EX_i - EX'_i)| \leq n^{-1} \sum_{i=1}^n E[|X_i|I(|X_i| > i)] \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Combining (3.8) and (3.9) we obtain

$$\lim_{n \rightarrow \infty} \frac{1}{n\varphi^{\frac{1}{2}}(n)} \sum_{i=1}^n (X_i - EX_i) = 0 \text{ a.s.}$$

□

**COROLLARY 3.4.** *If  $\omega(k) = 2^k$  (or  $\omega(k) = [q^k]$ ,  $q > 1$ ) and  $\{X_i, i \geq 1\}$  is  $\Delta_k$ -independent,  $P\{|X_i| \geq t\} \leq CP\{|X| \geq t\}$  for all nonnegative real numbers  $t$ ,  $E|X|^p < \infty$ , ( $1 \leq p < 2$ ), then*

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}}} \sum_{i=1}^n (X_i - EX_i) = 0 \text{ a.s.}$$

*Proof.* Really, in that case  $\varphi(i) = O(1)$ , so, from Corollary 3.3, we obtain

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{p}}} \sum_{i=1}^n (X_i - EX_i) = 0 \text{ a.s.}$$

□

## References

- [1] B. von Bahr and C. G. Esseen, *Inequalities for the  $r$ -th absolute moment of a sum of random variables*,  $1 \leq r \leq 2$ , Ann. Math. Statist. **36** (1965), 299–303.
- [2] B. D. Choi and S. H. Sung, *On convergence of  $(S_n - ES_n)/n^{1/r}$ ,  $1 < r < 2$ , for pairwise independent random variables*, Bull. Korean Math. Soc. **22** (1985), no. 2, 79–82.

- [3] N. Etemadi, *An elementary proof of the strong law of large numbers*, Z. Wahrsch. Verw. Gebiete **55** (1981), no. 1, 119–122.
- [4] V. F. Gaposhkin, *On the strong law of large numbers for blockwise-independent and blockwise-orthogonal random variables*, Theory Probab. Appl. **39** (1994), no. 4, 677–684.
- [5] ———, *Series of block-orthogonal and block-independent systems*, Izv. Vyssh. Uchebn. Zaved. Mat. (1990), no. 5, 12–18.
- [6] D. H. Hong and S. Y. Hwang, *Marcinkiewicz-type Strong law of large numbers for double arrays of pairwise independent random variables*, Int. J. Math. Math. Sci. **22** (1999), no. 1, 171–177.
- [7] D. H. Hong and A. I. Volodin, *Marcinkiewicz-type law of large numbers for double array*, J. Korean Math. Soc. **36** (1999), no. 6, 1133–1143.
- [8] F. Móricz, *Strong limit theorems for blockwise  $m$ -independent and blockwise quasi-orthogonal sequences of random variables*, Proc. Amer. Math. Soc. **101** (1987), no. 4, 709–715.

NGUYEN VAN QUANG, DEPARTMENT OF MATHEMATICS, UNIVERSITY OF VINH, 182  
LE DUAN, VINH, NGHEAN, VIETNAM  
*E-mail*: nvquang@hotmail.com

LE VAN THANH, DEPARTMENT OF MATHEMATICS, UNIVERSITY OF VINH, 182 LE  
DUAN, VINH, NGHEAN, VIETNAM  
*E-mail*: lvthanhvinh@yahoo.com