

BOUNDS ON PROBABILITY FOR THE OCCURRENCE OF EXACTLY r, t OUT OF m, n EVENTS

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ABSTRACT. Let A_1, A_2, \dots, A_m and B_1, B_2, \dots, B_n be two sequences of events on a given probability space. Let X_m and Y_n , respectively, be the number of those A_i and B_j , which occur we establish new upper and lower bounds on the probability $P(X = r, Y = t)$ which improve upper bounds and classical lower bounds in terms of the bivariate binomial moment $S_{r,t}, S_{r+1,t}, S_{r,t+1}$ and $S_{r+1,t+1}$.

1. Introduction

Let A_1, A_2, \dots, A_m and B_1, B_2, \dots, B_n be two sequences of events on the same probability space. Let $X = X_m(A)$ and $Y = Y_n(B)$, respectively, be the number of those A_i and B_j , which occur. Put $S_{o,o} = 1$ and for integers $r \geq 1$ and $t \geq 1$, set

$$(1) \quad S_{r,t} = \sum P(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_r} \cap B_{j_1} \cap B_{j_2} \cap \dots \cap B_{j_t})$$

where the summation \sum is over all subscripts satisfying $1 \leq i_1 < i_2 < \dots < i_r \leq m$ and $1 \leq j_1 < j_2 < \dots < j_t \leq n$. It can easily be proved, by turning to the method of indicators, that $S_{r,t}$, $1 \leq r \leq m, 1 \leq t \leq n$, is called the binomial moment of the vector (X, Y) , namely,

$$(2) \quad S_{r,t} = E \left[\binom{X}{r} \binom{Y}{t} \right].$$

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Bounds by linear combinations of the binomial moments $S_{i,j}$ on the distribution of the vector (X, Y) are called Bonferroni-Type Inequalities (B-T-I). There are two reasons for developing B-T-I. It can be utilized for obtaining limit theorems for normalized multivariate order statistics [see Galambos] and for actual computations of the bounds for fixed m and n . But the mentioned inequalities may become impractical for one of two reasons ; either (i) $S_{r,t}$ is known only for a very limited number of the values of r and t , or (ii) $S_{r,t}$ is so large that the successive terms, providing the upper and lower bounds, become trivial (Upper bounds exceed one and lower bounds become negative).

In this paper, we establish new upper and lower bounds on the distribution

$P(X = r, Y = t)$ which improve classical upper bounds and lower bounds in terms of $S_{r,t}$, $S_{r+1,t}$, $S_{r,t+1}$ and $S_{r+1,t+1}$.

2. The Results

Using the notation of the introduction, we shall prove the following results.

THEOREM 1. For integers r, t, m and n with $1 \leq r < m, 1 \leq t < n$,

$$(3) \quad P(X = r, Y = t) \geq \min(S_{r,t} - (r + 1)S_{r+1,t}, S_{r,t} - (t + 1)S_{r,t+1})$$

It is shown by Meyer that

$$(4) \quad P(X = r, Y = t) \geq S_{r,t} - (r + 1)S_{r+1,t} - (t + 1)S_{r,t+1}$$

Thus, it is obvious that inequality (3) is better than (4).

THEOREM 2. For integers r, t, m and n with $1 \leq r < m, 1 \leq t < n$,

$$(5) \quad P(X = r, Y = t) \geq S_{r,t} - (r + 1)S_{r+1,t} - (t + 1)S_{r,t+1} + \frac{(m + n - r - t - 1)}{(m - r)(n - t)}(r + 1)(t + 1)S_{r+1,t+1}.$$

Thus, it is obvious that inequality (5) is better than (4) if we know $S_{r+1,t+1}$.

THEOREM 3. For integers r, t, m and n with $1 \leq r < m, 1 \leq t < n$,

$$(6) \quad P(X = r, Y = t) \leq \min \left(S_{r,t} - \frac{r+1}{m-r} S_{r+1,t}, S_{r,t} - \frac{t+1}{n-t} S_{r,t+1} \right).$$

It is shown by Galambos and Lee(1994) that

$$(7) \quad P(X = r, Y = t) \leq S_{r,t} - \frac{(r+1)(t+1)}{(m-r)(n-t)} S_{r+1,t+1}.$$

Thus, it is obvious that inequality (6) is better than (7).

THEOREM 4. For integers r, t, m and n with $1 \leq r < m, 1 \leq t < n$,

$$(8) \quad \begin{aligned} P(X = r, Y = t) \leq & S_{r,t} - \frac{(r+1)}{(m-r)} S_{r+1,t} - \frac{(t+1)}{(n-t)} S_{r,t+1} \\ & + \frac{(r+1)(t+1)}{(m-r)(n-t)} S_{r+1,t+1}. \end{aligned}$$

This inequality is same as that of Galambos and Xu(1995).

3. proofs

PROOF OF THEOREM 1. We use the method of indicators. Let

$$I(X = r, Y = t) = I(X = r)I(Y = t) = \begin{cases} 1 & \text{if } X = r \text{ and } Y = t \\ 0 & \text{if otherwise} \end{cases}.$$

By using binomial moment of (2) and the method of indicators, the right hand side of (3) becomes

$$\begin{aligned} & \min E \left[\binom{X}{r} \binom{Y}{t} - (r+1) \binom{X}{r+1} \binom{Y}{t}, \binom{X}{r} \binom{Y}{t} \right. \\ & \quad \left. - (t+1) \binom{X}{r} \binom{Y}{t+1} \right] \\ & = \min E \left[\binom{X}{r} \binom{Y}{t} (1+r-X, 1+t-Y) \right] \end{aligned}$$

Since $E[I(X = r, Y = t)] = P(X = r, Y = t)$, and thus in order to prove (3), it suffices to show that

$$(9) \quad I(X = r)I(Y = t) \geq \min \left[\binom{X}{r} \binom{Y}{t} (1 + r - X, 1 + t - Y) \right]$$

Note that both sides of (9) are zero if either X or Y is less than r or t , respectively. Also, both sides of (9) are one if X and Y equal r and t , respectively. Thus, we have to prove that

(10) $f(X, Y) =$ (the right hand side of (9)) is non-negative for $X = r$ and $Y \geq t + 1$, $X \geq r + 1$ and $Y = t$, $X \geq r + 1$ and $Y \geq t + 1$.

(i) First case. For positive integers r, t with $X = r$ and $Y \geq t + 1$; that is, there are the events that exactly $r A_i$ and at least $t + 1 B_j$ which occur. Then

$$\begin{aligned} f(r, t + p) &= \min \left[\binom{t + p}{t} \binom{1, 1 + t - (t + p)}{1} \right] \\ &= \binom{t + p}{t} (1 - p) \leq 0 \text{ for integer } p \text{ with } 1 \leq p \leq n - t. \end{aligned}$$

(ii) Second case. For positive integers r, t with $X \geq r + 1$ and $Y = t$; that is, there are the event that exactly $t B_j$ and at least $r + 1 A_i$, which occur. Then

$$f(r + q, t) = \min \left[\binom{r + q}{r} \binom{1 + r - (r + q), 1}{1} \right] = \binom{r + q}{r} (1 - q) \leq 0$$

for integer q with $1 \leq q \leq n - r$.

(iii) Third case. For positive integers r, t, X and Y with $r + 1 \leq X$ and $t + 1 \leq Y$; that is, there are the events that at least $r + 1 A_i$ and at least $t + 1 B_j$ which occur. Then

$$f(r + q, t + p) = \min \left[\binom{r + q}{r} \binom{t + p}{t} \binom{1 + r - (r + p), 1 + t - (t + p)}{1} \right] \leq 0$$

for integers p, q with $1 \leq p \leq n - t, 1 \leq q \leq m - r$. Hence, we get (10). This completes the proof. \square

PROOF OF THEOREM 2. We use the method of indicators. Let

$$I(X = r, Y = t) = I(X = r)I(Y = t) = \begin{cases} 1 & \text{if } X = r \text{ and } Y = t \\ 0 & \text{if otherwise.} \end{cases}$$

By using binomial moment of (2) and the method of indicators, the right hand side of (5) becomes

$$\begin{aligned} & E \left[\binom{X}{r} \binom{Y}{t} - (r+1) \binom{X}{r+1} \binom{Y}{t} - (t+1) \binom{X}{r} \binom{Y}{t+1} \right. \\ & \quad \left. + \frac{(m+n-r-t-1)}{(m-r)(n-t)} (r+1)(t+1) \binom{X}{r+1} \binom{Y}{t+1} \right] \\ = & E \left[\binom{X}{r} \binom{Y}{t} [1+r+t-X-Y \right. \\ & \quad \left. + \frac{(m+n-r-t-1)}{(m-r)(n-t)} (X-r)(Y-t)] \right]. \end{aligned}$$

Then $E[I(X = r, Y = t)] = P(X = r, Y = t)$, and thus in order to prove (5), it suffices to show that

$$(11) \quad I(X = r)I(Y = t) \geq \binom{X}{r} \binom{Y}{t} \left[1+r+t-X-Y + \frac{(m+n-r-t-1)}{(m-r)(n-t)} (X-r)(Y-t) \right]$$

Note that both sides of (11) are zero if either X or Y is less than r or t, respectively. Also, both sides of (11) are one if X and Y equal r and t, respectively. Thus, we have to prove that

(12) $f(X, Y)$ = (the right hand side of (11) is non-negative) for $X = r$ and $Y \geq t + 1$, $X \geq r + 1$ and $Y = t$, $X \geq r + 1$ and $Y \geq t + 1$.

(i) First case. For positive integers r, t with $X = r$ and $Y \geq t + 1$; that is, there are exactly r A_i and at least $t + 1$ B_j which occur. Then

$$f(r, t + p) = \binom{t+p}{t} [1+t-(t+p)] \leq 0$$

for integer P with $1 \leq p \leq n - t$. Hence, we get (12).

(ii) Second case. For positive integers r, t with $X \geq r + 1$ and $Y = t$; that is, there are at least $r + 1$ A_i and exactly t B_j which occur. Then

$$f(r + q, t) = \binom{r + q}{r} [1 + r - (r + q)] \leq 0$$

for integer q with $1 \leq q \leq m - r$. Hence, we get (12).

(iii) Third case. For positive integers r, t, X and Y with $r + 1 \leq X$ and $t + 1 \leq Y$; that is, there are at least $r + 1$ A_i and at least $t + 1$ B_j which occur. Then

$$f(r + q, t + p) = \binom{r + q}{r} \binom{t + p}{t} \left[1 - q - p + \frac{(m + n - r - t - 1)}{(m - r)(n - t)} pq \right]$$

for integers p, q with $1 \leq p \leq n - t, 1 \leq q \leq m - r$. Let $g(q, p) = 1 - q - p + \frac{(m + n - r - t - 1)}{(m - r)(n - t)} pq$. Then $g(q, p), 1 \leq q \leq m - r, 1 \leq p \leq n - t$, attains its maximum at end point, yielding $g(m - r, n - t) = 0$. Thus, $f(r + q, t + p)$ is less than zero and equals zero. Hence, we get (12). This completes the proof. □

PROOF OF THEOREM 3. We have to prove that the following inequalities hold; that is,

$$(13) \quad P(X = r, Y = t) \leq S_{r,t} - \frac{r + 1}{m - r} S_{r+1,t} \text{ and}$$

$$(13') \quad P(X = r, Y = t) \leq S_{r,t} - \frac{t + 1}{n - t} S_{r,t+1}$$

By using the method of indicators and binomial moment of (2), they become

$$(14) \quad I(X = r)I(Y = t) \leq \binom{X}{r} \binom{Y}{t} - \frac{r + 1}{m - r} \binom{X}{r + 1} \binom{Y}{t} \text{ and}$$

$$(14') \quad I(X = r)I(Y = t) \leq \binom{X}{r} \binom{Y}{t} - \frac{t + 1}{n - t} \binom{X}{r} \binom{Y}{t + 1}$$

Thus, by proving (14) and (14') we obtain (13) and (13') by taking expectations. First, note that both (14) and (14') are valid if either X or Y is less than r or t , respectively, and both sides of (14) and (14') are one if X and Y equal r and t , respectively. Hence, for the sequel, we assume that $X \geq r + 1$ and $Y \geq t$, $X \geq r$ and $Y \geq t + 1$, $X \geq r + 1$ and $Y \geq t + 1$. Thus, (14) becomes

$$0 = I(X = r)I(Y = t) \leq \binom{X}{r} \binom{Y}{t} \left[1 - \frac{X - r}{m - r} \right]$$

Now, $h(X) = \binom{X}{r} \binom{Y}{t} \left[1 - \frac{X - r}{m - r} \right]$, $r \leq X \leq m$, attains its minimum at $x = m$, yielding $h(m) = 0$. Hence, we get (14).

The proof of (14') is identical to the above argument. The proof is completed. □

PROOF OF THEOREM 4. We also use the method of indicators.

$$\text{Let } I(X = r, Y = t) = I(X = r)I(Y = t) = \begin{cases} 1 & \text{if } X = r \text{ and } Y = t \\ 0 & \text{if otherwise.} \end{cases}$$

By using binomial moment of (2) and the method of indicators, the right hand side of (8) becomes

$$\begin{aligned} & E \left[\binom{X}{r} \binom{Y}{t} - \frac{(r + 1)}{(m - r)} \binom{X}{r + 1} \binom{Y}{t} - \frac{(t + 1)}{(n - t)} \binom{X}{r} \binom{Y}{t + 1} \right. \\ & \quad \left. + \frac{(r + 1)(t + 1)}{(m - r)(n - t)} \binom{X}{r + 1} \binom{Y}{t + 1} \right] \\ & = E \left[\binom{X}{r} \binom{Y}{t} \left(1 - \frac{(X - r)}{(m - r)} - \frac{(Y - t)}{(n - t)} + \frac{(X - r)(Y - t)}{(m - r)(n - t)} \right) \right] \end{aligned}$$

Then $E I(X = r, Y = t) = P(X = r, Y = t)$, and thus in order to prove (8), it suffices to show that

$$(15) \quad \begin{aligned} & I(X = r)I(Y = t) \\ & \leq \left[\binom{X}{r} \binom{Y}{t} \left(1 - \frac{(X - r)}{(m - r)} - \frac{(Y - t)}{(n - t)} + \frac{(X - r)(Y - t)}{(m - r)(n - t)} \right) \right]. \end{aligned}$$

Note that both sides of (15) are zero if either X or Y is less than r or t , respectively. Also, both sides of (15) are one if X and Y equal r and t , respectively. Thus, we have to prove that

(16) $f(X, Y) =$ the right hand side of (15) ≥ 0 for $X = r$ and $Y \geq t + 1$, $X \geq r + 1$ and $Y = t$, $X \geq r + 1$ and $Y \geq t + 1$.

(i) First case. For positive integers r, t with $X = r$ and $Y \geq t + 1$; that is, there are exactly r A_i and at least $t + 1$ B_j which occur. Then

$$f(r, t + p) = \binom{t + p}{t} \left[1 - \frac{p}{(n - t)} \right] \geq 0$$

for integer P with $1 \leq p \leq n - t$. Hence, we get (16).

(ii) Second case. For positive integers r, t with $X \geq r + 1$ and $Y = t$; that is, there are at least $r + 1$ A_i and exactly t B_j which occur. Then

$$f(r + q, t) = \binom{r + q}{r} \left[1 - \frac{q}{(m - r)} \right] \geq 0$$

for integer q with $1 \leq q \leq m - r$. Hence, we get (16).

(iii) Third case. For positive integers r, t, X and Y with $r + 1 \leq X$ and $t + 1 \leq Y$; that is, there are at least $r + 1$ A_i and at least $t + 1$ B_j which occur. Then

$$f(r + q, t + p) = \binom{r + q}{r} \binom{t + p}{t} \left[1 - \frac{q}{(m - r)} - \frac{p}{(n - t)} + \frac{qp}{(m - r)(n - t)} \right]$$

for integers p, q with $1 \leq p \leq n - t$, $1 \leq q \leq m - r$. Let

$$g(q, p) = \left[1 - \frac{q}{(m - r)} - \frac{p}{(n - t)} + \frac{qp}{(m - r)(n - t)} \right]$$

Then

$$g(q, p) = \frac{(m - r - q)(n - t - p)}{(m - r)(n - t)} \geq 0$$

for $1 \leq q \leq m - r$, $1 \leq p \leq n - t$. Thus, $f(r + q, t + p)$ is greater than zero and equals zero. Hence, we get (16). This completes the proof. \square

References

- [1] J. Galambos, *Order statistics of samples from multivariate distributions*, J. Amer. Stat. **70** (1975), 674-680.
- [2] J. Galambos, and M. -Y. Lee, *Further studies of Bonferroni-Type inequalities*, Journal of Applied probability Vol. 31A (1994), 63-69.
- [3] J. Galambos, and Y. Xu, *Bivariate Extension of the Method of Polynomials for Bonferroni-Type Inequalities*, Journal of Multivariate Analysis **52** (1995), 131-139.
- [4] R. M. Meyer, *Note on a multivariate form of Bonferroni's inequalities*, Ann, Math, Statist **40** (1969), 692-693.

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