# On the E-optimality of different blocksize designs

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### ABSTRACT

Constantine (1981) extended the results of Takeuchi (1961) by adding some new blocks to certain known E-optimal block designs. But they are confined to equal blocksize designs. In this paper we again generalize them to different blocksize case. By augmenting some known E-optimal block designs having blocks of equalsize with blocks of different sizes, additional E-optimal block designs are obtained.

## 1. Introduction

Let d be a block design with b blocks and v treatments. Then the information matrix of a block design d is defined as follows;

$$C(d) = \text{diag } [r_1, r_2, \dots, r_v] - N \text{ diag } [k_1^{-1}, k_2^{-1}, \dots, k_b^{-1}]N'$$
 (1·1)

where  $r_i$ 's are replicates of treatments,  $k_i$ 's are the sizes of blocks, and N is the incidence matrix whose entry  $n_{ij}$  gives the number of times treatment i occurs in block j.

We see that the matrix C(d) is symmetric, non-negative definite and has zero row sums. A design d is said to be connected if all contrasts of treatment effects are estimable. It is known that a design d with v treatments is connected iff rank (C(d)) = v-1. In this case we may assume the eigenvalues of C(d) are

$$0 = \mu_0(d) < \mu_1(d) \le \mu_2(d) \le \cdots \le \mu_{v-1}(d)$$

Let  $\Omega(v, b, k)$  denote the class of all connected block designs having v treatments

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<sup>†</sup> This research was supported by a grant from the Korea Science and Engineering Foundation, 1986.

arranged in b blocks of size k. A design  $d^*$  is said to be E-optimal in  $\Omega(v, b, k)$  if among all designs in  $\Omega(v, b, k)$ , it minimizes the supremum of the variance of least square estimators for normalized treatment effect contrasts.

With the aid of Ehrenfeld (1955), we have a criterion for E-optimality; A design  $d^*$   $\epsilon \Omega(v, b, k)$  is E-optimal in  $\Omega(v, b, k)$  iff

$$\mu_1(d^*) > \mu_1(d)$$

for any  $d \in \Omega(v, b, k)$ , where  $\mu_1(d)$  denotes the minimum positive eigenvalue of C(d).

If d is a combined design of designs  $d_1$  and  $d_2$  under the same treatments, then clearly

$$C(d) = C(d_1) + C(d),$$

and

$$\mu_1(d) \ge \mu_1(d_i), i=1,2$$
 (1.2)

Using the property (Rao(1973))

$$\mu_1(d) = \inf_{\xi} (\xi' C(d) \xi),$$

where  $\xi$  is a normalized  $v \times 1$  column vector with zero column sum, the following Lemmas can be shown in a similar way to Constantine (1981) and Jacroux (1982).

**Lemma 1**: Let  $C(d) = (c_{ij}(d))$  be the  $v \times v$  information matrix of design d. Then

$$\mu_1(d) \leq \frac{c_{ii}(d) + c_{ij}(d) - 2c_{ij}(d)}{2}, i \neq j.$$

**Lemma 2**: Let C(d) be as in Lemma 1, and M be a proper subset of the treatments  $\{1, 2, \dots, v\}$ , say  $M = \{1, 2, \dots, m\}$ ,  $1 \le m < v$ . Then

$$\mu_1(d) \leq \frac{v}{m(v-m)} \sum_{i,j \in M} c_{ij}(d)$$

### 2. Main Retsults

Readers may refer to Raghavarao(1971) for the usual terminologies and notations on block design throughout the paper. Let  $\Omega(v; b_1, b_2; k_1, k_2)$  be the class of all connected block designs having v treatments arranged in  $b_i$  blocks of size  $k_i$  for i=1,2.

The two theorems given in this section are extensions of the results proven in Constantine (1981) and Jacroux (1982).

**Theorem 1**: Let  $d_1^* \in \Omega(v, b_1, k_1)$  be a BIB design and  $d_2^* \in \Omega(v, b_2, k_2)$  be an arbitrary binary design based on the same treatments with  $d_1^*$ . If  $k_1 \ge k_2$  and  $b_2 k_2 < v$ , then the combined design  $d^* = (d_1^*, d_2^*)$  is E-optimal in  $\Omega(v; b_1, b_2; k_1, k_2)$ .

**Proof**: Since  $d_1^*$  is a BIB design,

$$r(d_1^*) = b_1 k_1 / v$$
 and  $\lambda(d_1^*) = b_1 k_1 (k_1 - 1) / (v(v - 1))$ 

and we can see that

$$\mu_1(d_1^*) = \frac{r(d_1^*)(k_1 - 1) + \lambda(d_1^*)}{k_1}$$

So by  $(1\cdot 2)$ , it follows that

$$\mu_1(d^*) \ge \frac{b_1(k_1-1)}{v-1}$$
 (2·1)

Now let d be any design in  $\Omega(v; b_1, b_2; k_1, k_2)$ . Then d can be represented by  $d = (d_1, d_2)$  for some  $d_1 \in \Omega(v, b_1, k_1)$  and  $d_2 \in \Omega(v, b_2, k_2)$ . Hence  $C(d) = C(d_1) + C(d_2)$ , and so

$$r_i(d) = r_i(d_1) + r_i(d_2), i=1,2,\dots,v$$

and  $\lambda_{ij}(d) = \lambda_{ij}(d_1) + \lambda_{ij}(d_2), i \neq j.$ 

If we let  $r_{i_0}(d) = \min_i r_i(d)$ , then by the condition  $\frac{b_2 k_2}{v} < 1$  we get  $r_{i_0}(d) \le r(d_1^*)$  since  $r_i(d)$  is the replication of the *i*-th treatment. Then from Lemma 2 with m=1 and the condition  $k_1 \ge k_2$ , it follows that

$$\mu_1(d) \le \frac{b_1(k_1 - 1)}{v - 1}$$
 (2·2)

The result now follows from  $(2 \cdot 1)$  and  $(2 \cdot 2)$ .

Theorem 2: Let  $d_1^*$  be a group divisible designs with parameters v=mn(m) groups containing n treatments each),  $b_1, k_1, \lambda_1, \lambda_2 = \lambda_1 + 1$ , and  $d_2^* \in \Omega$   $(v, b_2, k_2)$  be an arbitrary design based on the same treatments with  $d_1^*$ . If  $k_1 \ge k_2$  and  $b_2 k_2 < v-m$ , then the combined design  $d^* = (d_1^*, d_2^*)$  is E-optimal in  $\Omega$   $(v; b_1, b_2; k_1, k_2)$ .

**Proof**: Since  $d_1^*$  is a group divisible design,

$$r(d_1^*) = \frac{b_1 k_1}{v}, \quad \lambda_1(d_1^*) = \left[\frac{b_1 k_1 (k_1 - 1)}{v(v - 1)}\right]$$

where [ ] denotes the integer part of the number, and

$$\mu_1(d_1^*) = \frac{r(d_1^*)(k_1-1) + \lambda_1(d_1^*)}{k_1}$$

Since  $b_2k_2 < v-m$ , there exist at least 2 treatments i, j in  $d^*$  such that

$$c_{ii}(d^*) = c_{jj}(d^*) = \frac{r(d_1^*)(k_1-1)}{k_1}, \ c_{ij}(d^*) = -\frac{\lambda_1(d_1^*)}{k_1}$$

from (1, 1).

Then by Lemma 1,

$$\mu_1(d^*) \leq \frac{r(d_1^*)(k_1-1) + \lambda_1(d_1^*)}{k_1}$$

Since  $\mu_1(d^*) \ge \mu_1(d_1^*)$  by (1·2), it follows that

$$\mu_1(d^*) = \frac{r(d_1^*)(k_1-1) + \lambda_1(d_1^*)}{k_1}$$

Now let  $d \in \Omega(v; b_1, b_2; k_1, k_2)$  be arbitrary. Then  $d = (d_1, d_2)$  for some  $d_1 \in \Omega(v, b_1, k_1)$  and  $d_2 \in \Omega(v, b_2, k_2)$ .

Here we consider the two cases as follows;

- (i) there exists some  $r_1(d)$  such that  $r_i(d) < r(d_1^*)$ ,
- (ii)  $r_i(d) \ge r(d_1^*)$  for  $i = 1, 2, \dots, v$ .

In case (i), by the condition  $k_1 \ge k_2$  and  $c_{ii}(d) = c_{ii}(d_1) + c_{ii}(d_2)$  we can show

$$c_{ii}(d) \leq \frac{(k_1-1)(r(d_1^*)-1)}{b_1}$$

Then by Lemma 2 and some algebra works analogous to those of Constantine (1981), we have

$$\mu_1(d) \leq \frac{v}{v-1} \frac{(k_1-1)(r(d_1^*)-1)}{k_1} \leq \frac{r(d_1^*)(k_1-1)+\lambda_1(d_1^*)}{k_1}$$

Therefore  $\mu_1(d) \leq \mu_1(d^*)$  in this case.

In a similar way, case (ii) is proved.

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