

## Complete convergence for weighted sums of AANA random variables

TAE-SUNG KIM<sup>1</sup> AND MI-HWA KO<sup>2</sup>

### Abstract

We study maximal second moment inequality and derive complete convergence for weighted sums of asymptotically almost negatively associated(AANA) random variables by applying this inequality.

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### 1. Introduction

Recall that a finite family  $\{X_i, 1 \leq i \leq n\}$  is said to be negatively associated(NA) if for any disjoint subsets  $A, B \subset \{1, \dots, n\}$  and any real coordinatewise nondecreasing functions  $f : R^A \rightarrow R$  and  $g : R^B \rightarrow R$ ,  $Cov(f(X_i, i \in A), g(X_j, j \in B)) \leq 0$  (see Joag-Dev and Proschan(1983)). Matula(1992) has established a maximal inequality for negatively associated(NA) sequences. By inspecting the proof of Matula's(1992) maximal inequality, Chandra and Ghosal(1996) found that one can also allow positive correlations provided they are small. A sequence  $\{X_n, n \geq 1\}$  of random variables is called asymptotically almost negatively associated(AANA) if there is a nonnegative sequence  $q(m) \rightarrow 0$  such that

$$\begin{aligned} &Cov(f(X_m), g(X_{m+1}, \dots, X_{m+k})) \\ &\leq q(m)(var(f(X_m))var(g(X_{m+1}, \dots, X_{m+k})))^{1/2} \end{aligned} \quad (1)$$

for all  $m, k \geq 1$  and for all coordinatewise increasing continuous functions  $f$  and  $g$  whenever the right side of (1) is finite.

In this paper we study complete convergence for weighted sums of AANA sequence, which has never been studied previously in the literature.

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<sup>1</sup>Professor, Department of Mathematics, WonKwang University Iksan, Jeonbuk 570-749  
E-mail : starkim@wonkwang.ac.kr

<sup>2</sup>Instructor, Department of Mathematics, WonKwang University Iksan, Jeonbuk 570-749  
E-mail : songhack@wonkwang.ac.kr

## 2. Results

**Lemma 2.1** Let  $\{X_n, n \geq 1\}$  be a sequence of asymptotically almost negatively associated(AANA). Then  $\{f_n(X_n), n \geq 1\}$  is still a sequence of AANA random variables, where  $f_n(\cdot), n = 1, 2, \dots$ , are nondecreasing functions.

**Lemma 2.2 (Chandra and Ghosal (1996))** Let  $X_1, \dots, X_n$  be mean zero, square integrable random variables such that (1) holds for  $1 \leq m < k + m \leq n$  and for all coordinatewise nondecreasing continuous functions  $f$  and  $g$  whenever the right-hand side of (1) is finite. Let  $A^2 = \sum_{m=1}^{n-1} q^2(m)$  and  $\sigma_k^2 = EX_k^2, 1 \leq k \leq n$ . Then we have

$$E\left(\max_{1 \leq k \leq n} \sum_{i=1}^k X_i\right)^2 \leq (A + (1 + A^2)^{1/2})^2 \sum_{k=1}^n \sigma_k^2. \quad (2)$$

**Theorem 2.3** Let  $\{X_k, k \geq 1\}$  be a sequence of AANA random variables with  $EX_k = 0$  and  $EX_k^2 < \infty$  for all  $k \geq 1$ . Let  $A^2 = \sum_{m=1}^{n-1} q^2(m)$  and let  $\{a_{nk}, 1 \leq k \leq n, n \geq 1\}$  be an array of real numbers satisfying the condition

$$\sum_{k=1}^n a_{nk}^2 = O(n^\delta) \text{ as } n \rightarrow \infty \text{ for some } 0 < \delta < 1. \quad (3)$$

Assume that

$$\sum_{m=1}^{\infty} q^2(m) < \infty. \quad (4)$$

Then,  $\forall \epsilon > 0$  and  $\delta'$  such that  $\delta < \delta' \leq 1$

$$\sum_{n=1}^{\infty} n^{-\delta'} P\left(\max_{1 \leq k \leq n} \left| \sum_{i=1}^k a_{ni} X_i \right| > \epsilon n^{1/2}\right) < \infty. \quad (5)$$

**Proof .** To prove (5) it suffices to show that

$$\sum_{n=1}^{\infty} n^{-\delta'} P\left(\max_{1 \leq k \leq n} \left| \sum_{i=1}^k a_{ni}^+ X_i \right| > \epsilon n^{1/2}\right) < \infty \quad \forall \epsilon > 0, \quad (6)$$

$$\sum_{n=1}^{\infty} n^{-\delta'} P\left(\max_{1 \leq k \leq n} \left| \sum_{i=1}^k a_{ni}^- X_i \right| > \epsilon n^{1/2}\right) < \infty \quad \forall \epsilon > 0, \quad (7)$$

where  $a_{ni}^+ = a_{ni} \vee 0, a_{ni}^- = (-a_{ni}) \vee 0$ . We need only to prove (6), since the proof of (7) is analogous. From Lemma 2.1  $\{a_{ni}^+ X_i, 1 \leq i \leq n, n \geq 1\}$  is an AANA sequence, and

hence by applying we have Lemma 2.2

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{-\delta'} P(\max_{1 \leq k \leq n} |\sum_{i=1}^n a_{ni}^+ X_i| > \epsilon n^{1/2}) \\ & \leq \epsilon^{-2} \sum_{n=1}^{\infty} n^{-1-\delta'} (A + (1 + A^2)^{1/2})^2 \sum_{k=1}^n a_{nk}^2 EX_k^2 =: I. \end{aligned}$$

Note that conditions (3) and  $EX_k^2 < \infty$  imply

$$\sum_{k=1}^n a_{nk}^2 EX_k^2 = O(n^\delta) \text{ as } n \rightarrow \infty. \quad (8)$$

Hence, by (4) and (8) we have

$$I \ll (A + (1 + A^2)^{1/2})^2 \sum_{n=1}^{\infty} n^{-(1+\delta'-\delta)} < \infty \text{ for } 0 < \delta < \delta' \leq 1,$$

where  $a \ll b$  means  $a = O(b)$ . The proof is completed.

**Theorem 2.4** Let  $\{X, X_k, k \geq 1\}$  be sequence of identically distributed AANA random variables with  $EX = 0$ ,  $EX^2 < \infty$ . Let  $A^2 = \sum_{m=1}^{n-1} q^2(m)$  and let  $\{a_{nk}, 1 \leq k \leq n, n \geq 1\}$  be an array of real numbers satisfying (3). If (4) and (5) hold then for some  $\delta'$  such that  $\delta < \delta' \leq 1$

$$\sum_{n=1}^{\infty} n^{-\delta'} \sum_{j=1}^n P(|a_{nj} X_j| > n^{1/2}) < \infty. \quad (9)$$

**Proof.** From (5) we obviously get

$$\sum_{n=1}^{\infty} n^{-\delta'} P(\max_{1 \leq j \leq n} |a_{nj} X_j| > n^{1/2}) < \infty, \quad (10)$$

$$P(\max_{1 \leq j \leq n} |a_{nj} X_j| > n^{1/2}) \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (11)$$

Note that

$$P(\max_{1 \leq j \leq n} |a_{nj} X_j| > n^{1/2}) = \sum_{j=1}^n P(|a_{nj} X_j| > n^{1/2}, \max_{1 \leq i \leq j-1} |a_{ni} X_i| \leq n^{1/2}).$$

Hence, we deduce that

$$\begin{aligned} & \sum_{j=1}^n P(|a_{nj} X_j| > n^{1/2}) = P(\max_{1 \leq j \leq n} |a_{nj} X_j| > n^{1/2}) \\ & + \sum_{j=1}^n P(|a_{nj} X_j| > n^{1/2}, \max_{1 \leq i \leq j-1} |a_{ni} X_i| > n^{1/2}). \end{aligned} \quad (12)$$

Also, we have

$$\begin{aligned}
 & \sum_{j=1}^n P(|a_{nj}X_j| > n^{1/2}, \max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2}) \\
 &= \sum_{j=1}^n \{E[I(|a_{nj}X_j| > n^{1/2})I(\max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2})] \\
 & \quad - EI(|a_{nj}X_j| > n^{1/2})EI(\max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2})\} \\
 & \quad + \sum_{j=1}^n \{EI(|a_{nj}X_j| > n^{1/2})EI(\max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2})\} \\
 & \leq E \sum_{j=1}^n [I(|a_{nj}X_j| > n^{1/2}) - P(|a_{nj}X_j| > n^{1/2})]I(\max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2}) \\
 & \quad + \sum_{j=1}^n P(|a_{nj}X_j| > n^{1/2})P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}) = II + III. \tag{13}
 \end{aligned}$$

Define

$$Y_{nj} = \begin{cases} a_{nj}X_j, & \text{if } a_{nj} \geq 0, \\ -a_{nj}X_j, & \text{if } a_{nj} < 0. \end{cases}$$

Then  $\{Y_{nj}, 1 \leq j \leq n, n \geq 1\}$  and  $\{I(Y_{nj} > n^{1/2})\}$  are AANA by Lemma 2.1. By the Cauchy-Schwarz inequality and Lemma 2.2,

$$\begin{aligned}
 |II| &= |E \sum_{j=1}^n [I(|a_{nj}X_j| > n^{1/2}) - P(|a_{nj}X_j| > n^{1/2})] \\
 & \quad \times I(\max_{1 \leq i \leq j-1} |a_{ni}X_i| > n^{1/2})| \\
 & \leq [E(\sum_{j=1}^n I(|a_{nj}X_j| > n^{1/2}) - EI(|a_{nj}X_j| > n^{1/2}))^2 \\
 & \quad \times E(I(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}))^2]^{1/2} \\
 & = [\text{Var}(\sum_{j=1}^n I(|a_{nj}X_j| > n^{1/2}))P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2})]^{1/2} \\
 & \leq [2\{\text{Var}[\sum_{j=1}^n I(Y_{nj} > n^{1/2})] + \text{Var}[\sum_{j=1}^n I(Y_{nj} < -n^{1/2})]\} \\
 & \quad \times P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2})]^{1/2} \\
 & \leq [8\{\sum_{j=1}^n P(Y_{nj} > n^{1/2}) + \sum_{j=1}^n P(Y_{nj} < -n^{1/2})\} \\
 & \quad \times P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2})(A + (1 + A^2)^{1/2})^2]^{1/2}
 \end{aligned}$$

$$\begin{aligned} &\leq \frac{1}{2} \sum_{j=1}^n P(|a_{nj}X| > n^{1/2}) + 4(A + (1 + A^2)^{1/2})^2 \\ &\quad \times P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}) \end{aligned} \quad (14)$$

by  $\sqrt{ab} \leq \frac{a+b}{2}$ . From (12)-(14) we have

$$\begin{aligned} &\frac{1}{2} \sum_{j=1}^n P(|a_{nj}X| > n^{1/2}) \\ &\leq \{1 + 4(A + (1 + A^2)^{1/2})^2\} P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}) \\ &\quad + \sum_{j=1}^n P(|a_{nj}X| > n^{1/2}) P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}) \end{aligned} \quad (15)$$

and from (11) we get

$$\sum_{j=1}^n P(|a_{nj}X| > n^{1/2}) \ll P(\max_{1 \leq i \leq n} |a_{ni}X_i| > n^{1/2}), \quad (16)$$

for sufficiently large  $n$ . Therefore, from (10) and (16) the desired result (9) follows.

## References

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